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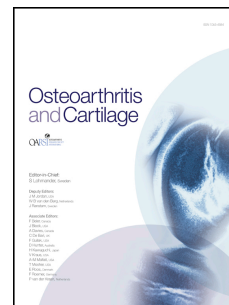
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Cardiometabolic traits mediating the effect of education on osteoarthritis risk: a Mendelian randomization study

Education, cardiometabolic traits and OA

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1 Abstract

2 Objective

3 To investigate which cardiometabolic factors underlie clustering of osteoarthritis with cardiovascular
4 disease, and the extent to which these mediate an effect of education.

5 Design

6 Genome-wide association study (GWAS) of osteoarthritis was performed in UK Biobank (60,800
7 cases and 328,251 controls) to obtain genetic association estimates for osteoarthritis risk. Genetic
8 instruments and association estimates for body mass index (BMI), low-density lipoprotein
9 cholesterol (LDL-C), systolic blood pressure (SBP), smoking and education were obtained from
10 existing GWAS summary data (sample sizes 188,577-866,834 individuals). Two-sample Mendelian
11 randomization (MR) analyses were performed to investigate the effects of exposure traits on
12 osteoarthritis risk. MR mediation analyses were undertaken to investigate whether the
13 cardiometabolic traits mediate any effect of education on osteoarthritis risk.

14 Results

15 MR analyses identified protective effects of higher genetically predicted education (main MR
16 analysis odds ratio [OR] per standard deviation increase 0.59, 95% confidence interval [CI] 0.54-0.64)
17 and LDL-C levels (OR 0.94, 95%CI 0.91-0.98) on osteoarthritis risk, and unfavourable effects of higher
18 genetically predicted BMI (OR 1.82, 95%CI 1.73-1.92) and smoking (OR 2.23, 95%CI 1.85-2.68). There
19 was no strong evidence of an effect of genetically predicted SBP on osteoarthritis risk (OR 0.98, 95%
20 CI 0.90-1.06). The proportion of the effect of genetically predicted education mediated through
21 genetically predicted BMI and smoking was 35% (95%CI 13%-57%).

1 **Conclusions**

2 These findings highlight education, obesity and smoking as common mechanisms underlying
3 osteoarthritis and cardiovascular disease. These risk factors represent clinical and public health
4 targets for reducing multi-morbidity related to the burden these common conditions.

5

6 **Key words**

7 Osteoarthritis, education, cardiometabolic, Mendelian randomization

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1 Introduction

2 Osteoarthritis (OA) is the most common form of arthritis worldwide and there are currently no
3 disease-modifying agents available. It accounts for 2.4% of all years lived with disability (YLD) and
4 ranks as a leading contributor to global YLDs (1). The prevalence of hip and knee OA worldwide is
5 close to 5% and is expected to increase further (1). Recent research on modifiable risk factors for OA
6 have investigated the influence of cardiovascular disease and educational level. An increased
7 prevalence of cardiovascular disease is found in OA (2). It is also recognised that lower educational
8 level is associated with increased cardiovascular disease (3, 4). However, the underlying mechanisms
9 are not well understood. The effect of education on OA risk may in part be mediated cardiovascular
10 risk factors that increase OA risk (3, 4), and evaluation of these risk factors could help to optimise
11 disease prevention at a clinical and public health level.

12 Assessing causal effects in observational research is difficult due to environmental confounding or
13 reverse causation. The Mendelian randomization (MR) approach can overcome some of these
14 limitations by using genetic variants related to the exposure of interest as instrumental variables for
15 investigating its effects on an outcome (5). Genetic variants are randomly allocated at conception,
16 and therefore their associations with the outcome are less affected by environmental confounding.
17 More recently, MR methods have been applied to investigate mediating pathways (6, 7), where use
18 of genetic variants that capture lifetime exposure also help overcome bias related to measurement
19 error that can hinder observational research.

20 The aim of this study was to apply the MR framework to investigate the effects of education and
21 cardiometabolic risk factors on risk of OA. For cardiometabolic risk factors that the MR analyses
22 supported to have a causal effect on OA risk, we aimed to further apply MR mediation analyses to
23 investigate the degree to which these factors might be mediating the effects of educational
24 attainment.

1 Methods

2 Overall study design

3 A genome-wide association study (GWAS) for OA was performed in the UK Biobank to obtain genetic
4 association estimates for OA risk. UK Biobank identification of OA was based on hospital-diagnosed
5 OA cases (8, 9). Genetic association estimates for cardiometabolic cardiovascular risk factors and
6 educational attainment (referred to hereafter as education) were selected from published GWASs.
7 MR analyses were performed to investigate the effects of their respective genetically predicted
8 levels on OA risk. The considered cardiometabolic risk factors were body mass index (BMI), low-
9 density lipoprotein cholesterol (LDL-C), lifetime smoking (referred to hereafter as smoking) and
10 systolic blood pressure (SBP). For cardiometabolic risk factors for which there was MR evidence of a
11 detrimental effect of their genetically predicted levels on OA risk, MR mediation analyses were
12 performed to investigate the degree to which this mediated any effect of education on OA risk.

13 Osteoarthritis genome-wide association study

14 GWAS for OA was performed in the UK Biobank (8, 9), a prospective cohort study of approximately
15 half a million participants with linked self-reported outcomes, health care records and genetic data.
16 OA cases in UK Biobank were defined based on International Classification of Diseases (ICD)-9 coding
17 (715, 721.0-721.42), ICD-10 coding (M15-M19, M47), and Office of Population Censuses and Surveys
18 Classification of Interventions and Procedures version 4 (OPCS-4) coding (W37-W42, W52-W54,
19 W58, W93-W95) and self-report (Supplementary Table 1 and Supplementary Figure 1). For GWAS
20 analysis, only white British participants defined by the UK Biobank genotyping quality control (QC)
21 were included (8). Baseline characteristics are provided in Table 1. A total of 60,800 OA cases and
22 328,251 controls were included in the primary analysis. For GWAS, we used REGENIE (10), a ridge
23 regression based method using Firth fallback regression correcting for age, sex, the first 20 genetic
24 principal components, genotyping chip and assessment center. In a sensitivity analysis, we also
25 included 19,846 OA cases which were of self-report only.

1 Exposure genetic association estimates

2 Genetic association estimates for BMI were obtained from the GIANT Consortium GWAS meta-
3 analysis of 806,834 European-ancestry individuals (11). Genetic association estimates for fasting LDL-
4 C were obtained from the Global Lipids Genetic Consortium GWAS of 188,577 European-ancestry
5 individuals that were not taking lipid lowering medication (12). Genetic association estimates for SBP
6 were obtained from a GWAS of 318,417 white British individuals performed in the UK Biobank. The
7 mean SBP from two automated recordings taken two minutes apart at baseline assessment were
8 used, and correction for any (self-reported) anti-hypertensive medication use was made by adding
9 10mmHg (3). Genetic association estimates for smoking were obtained from a GWAS of 462,690
10 European-ancestry individuals in the UK Biobank (13). A continuous lifetime smoking measure was
11 constructed from self-reported age at initiation, age at cessation and cigarettes smoked per day (13).
12 Genetic association estimates for education were obtained from a GWAS of 766,345 European-
13 ancestry individuals (14). Education was measured as the number of years completed in full time
14 education and was matched across different cohorts using the International Standard Classification
15 of Education system (15). Full details of GWAS analyses are available in their original publications.
16 We obtain genetic association estimates for education, BMI, SBP and smoking from studies that
17 included UK Biobank participants (3, 11, 13, 14), with OA genetic association estimates also obtained
18 from an overlapping UK Biobank population. Such participant overlap can result in bias of MR
19 estimates towards the observational estimate in the context of weak instruments (16). For
20 sensitivity analysis, we conducted the analysis for education and BMI using GWAS summary statistics
21 from non-overlapping populations (15, 17). For SBP and smoking, if evidence for association in MR
22 was found, we estimated the potential bias due to sample overlap as previously described (16).

23 Instrument selection

24 Instruments for each considered exposure in univariable MR analyses were selected as single-
25 nucleotide polymorphisms (SNPs) that associated with the exposure at genome-wide significance

1 ($P < 5 \times 10^{-8}$) and were independent, i.e. pairwise linkage disequilibrium (LD) $r^2 < 0.001$. To select
2 instruments for multivariable MR (MVMR) in analyses investigating mediators of the effect of
3 genetically predicted education on OA risk, all SNPs related to education or investigated mediators
4 at genome-wide significance were pooled and clumped to pairwise LD $r^2 < 0.001$ based on the lowest
5 P -value for their association with any trait. All clumping was performed using the TwoSampleMR
6 package in R (18).

7 Genetic association estimates for different traits were aligned to correspond to the same effect
8 allele. Palindromic variants were excluded in the main analysis and included for sensitivity analysis.
9 Only genetic variants for which association estimates were present for all traits being studied in a
10 given analysis were considered, and proxies were not used.

11 To quantify the ability to detect putative causal associations based on the available summary
12 statistics, we calculated the minimum detectable odds ratios (OR) for the risk of OA in MR analysis of
13 each exposure separately, given 80% power, type I error rate = 0.05, exposure GWAS summary
14 statistics sample size and the total variance explained by the genetic instruments (19). To evaluate
15 instrument strength, F statistics were calculated for individual genetic instruments.

16 Univariable Mendelian randomization

17 Multiplicative random-effects inverse-variance weighted (IVW) MR was used as the main analysis for
18 estimating the effects of genetically predicted cardiovascular risk factors and education on OA risk
19 (20). The genetic association estimates for the OA risk were the coefficients from logistic regression
20 (i.e. log odds ratios) for each genetic variant. The resulting MR estimate was exponentiated to obtain
21 the OR estimate given by MR.

22 When using multiple genetic variants as instrumental variables in MR, a potential source of bias is
23 horizontal pleiotropy, where the genetic variants affect the risk of OA through pathways
24 independent of the considered exposure. To assess the robustness of the findings to the potential
25 bias due to horizontal pleiotropy, we used contamination-mixture method, MR-Egger and weighted

1 median MR as sensitivity analyses (21-23). The contamination-mixture model assumes that MR
2 estimates from valid instruments follow a normal distribution centered on the true causal effect
3 estimate and that those calculated from invalid instrument variants follow a normal distribution with
4 their effect estimates centered on zero (22). A likelihood function is then maximized for allocating
5 each variant to one of the two mixture distributions (22). MR-Egger performs a regression of the
6 variant-outcome genetic association estimates on the variant-exposure genetic association
7 estimates, weighted for the precision of the variant-outcome genetic association estimates (23). The
8 slope of the regression line represents the MR estimate, and evidence for directional pleiotropy can
9 be evaluated by testing whether the intercept differs from zero (23). In weighted median MR, the
10 MR estimates from individual variants are ordered by their magnitude weighted for their precision,
11 and the median is selected as the overall MR estimate, with standard errors calculated by
12 bootstrapping (21). The MendelianRandomization package of R was used for performing all these
13 univariable MR analyses (24). The discrepancy between the main IVW MR analysis and sensitivity
14 analysis was used to assess for the potential presence of bias related to pleiotropic variants.

15 All MR estimates were calculated per one standard deviation (SD) unit increase in the exposure
16 under consideration, with SD estimates derived from UK Biobank data. For BMI this was 4.77kg/m^2 ,
17 for LDL-C this was 0.87mmol/l , for SBP this was 18.68mmHg and for education this was 3.6 years. For
18 smoking, a one standard deviation increase was equivalent to an individual smoking 20 cigarettes
19 per day for 15 years and stopping 17 years ago, for example (13).

20 **Multivariable Mendelian randomization**

21 The genetically predicted cardiovascular risk factors that showed evidence for a detrimental effect
22 on the risk of OA in univariable MR were taken forward for MVMR mediation analysis (7, 27). We
23 aimed to estimate the degree to which the effect of education on the risk of OA is mediated by the
24 cardiovascular risk factors.

1 In MVMR, the total effect of each exposure is decomposed to direct and indirect effects. This allows
2 for estimation of potential mediating effects and the proportion of the effect of the main exposure
3 of interest on the outcome that acts via other considered exposures (28, 29). Specifically, variant-OA
4 genetic association estimates (on the log odds ratio scale) were regressed on variant-education and
5 variant-cardiovascular risk factor genetic association estimates, weighted for the precision (i.e. the
6 inverse of their variance) of the variant-OA genetic association estimates and with the intercept
7 fixed at zero (27). The considered cardiovascular risk factors were included in this model both
8 individually and all together. The final OR estimate of the effect of education on the risk of OA from
9 MVMR was obtained by exponentiating the corresponding effect estimate. To estimate the
10 proportion of the effect of genetically predicted education on OA risk that was mediated through the
11 considered cardiovascular risk factors, the MR estimate for the effect of genetically predicted
12 education on OA risk after adjusting for genetically predicted levels of the cardiovascular risk factors
13 was divided by the total effect of education on OA risk estimated in the IVW univariable MR and
14 subtracted from 1, with standard errors estimated using the propagation of error method (6, 7).

15 **Measuring the strength of evidence**

16 No formal *P* value threshold for statistical significance was used. Instead, we interpret the evidence
17 provided by the results by looking at the effect size of interest and the width of its confidence
18 interval, combined with the consistency of the results across the different methods used(30).

19 **Ethical approval, data availability and reporting**

20 All data used in this work are publicly available and the studies from which they were obtained had
21 previously obtained relevant ethical approval and participant consent (8, 9, 11, 12). All data and
22 results generated in this work are presented in the main manuscript and the related supplementary
23 files. The reporting of this MR study follows the recommendations of the STROBE-MR Guidelines
24 (31), as detailed in the Supplementary Checklist. The codes for analysis are available from the
25 authors upon request.

1 Results

2 All genetic association estimates and their F statistics used in the univariable and multivariable MR
3 analyses are provided in Supplementary Tables 2-9 and visualized in Supplementary Figures 2-9. The
4 minimum detectable ORs on the risk of OA for each outcome are given in Supplementary Table 10.

5 In the univariable MR, there was evidence of a protective effect of genetically predicted education
6 and LDL-C on OA risk in the main IVW analyses (education: OR 0.59, 95% confidence interval [CI]
7 0.54-0.64; LDL-C: OR 0.94, 95%CI 0.91-0.98), with consistent findings in sensitivity analyses (Figure
8 1). There was evidence of an unfavourable effect of genetically predicted BMI and smoking on OA
9 risk in the main IVW MR analyses (BMI: OR 1.82, 95%CI 1.73-1.92; smoking: OR 2.23, 95%CI 1.85-
10 2.68), with consistent results obtained in sensitivity analyses (Figure 1). Similar results were obtained
11 when using non-overlapping summary statistics for BMI and education (Supplementary Table 11;
12 Supplementary Figure 10). The bias due to sample overlap in the log odds ratio for smoking and the
13 risk of OA under the null hypothesis was estimated at 0.012 and the expected Type I error rate for a
14 two-sided test with $\alpha = 0.05$ was estimated at 0.053. There was no evidence of an effect of
15 genetically predicted SBP on OA risk in the main IVW (OR 0.98, 95%CI 0.90-1.06) or any MR
16 sensitivity analysis (Figure 1). The MR-Egger intercept tests did not give evidence for the presence of
17 directional pleiotropy for education ($P=0.79$), LDL-C: ($P=0.21$), and SBP ($P=0.25$). There was weak
18 evidence for directional pleiotropy for BMI ($P=0.10$) and smoking: ($P=0.09$), however in both cases
19 MR-Egger estimate was consistent with the IVW estimate (Figure 1, Supplementary Table 11). Given
20 the identified effects of higher genetically predicted BMI and higher genetically predicted smoking
21 on increasing OA risk, multivariable MR mediation analyses were performed to investigate the
22 degree to which these traits were mediating the effect of genetically predicted education on OA risk.

23 The protective effect of genetically predicted education on OA risk attenuated from OR of 0.59
24 (95%CI 0.54 - 0.64) in IVW univariable analysis to OR of 0.66 (95%CI 0.60 - 0.73) after adjusting for
25 genetically predicted BMI in MVMR analysis, to OR of 0.67 (95%CI 0.61 - 0.74) after adjusting for

1 genetically predicted smoking in MVMR analysis, and to OR of 0.71 (95%CI 0.64 - 0.79) after
2 adjusting for both genetically predicted BMI and genetically predicted smoking in MVMR analysis
3 (Figure 2).

4 The proportion of the effect of genetically predicted education mediated through genetically
5 predicted BMI, smoking, and both BMI and smoking together was estimated as 23% (95%CI 1%-
6 44%), 25% (95%CI -3%-47%) and 35% (95%CI 13%-57%), respectively (Figure 3). The results obtained
7 by using genetic variant estimates from non-overlapping data sources showed similar directions of
8 the mediated proportions, albeit with higher uncertainty in the estimates (Supplementary Table 12).

9 Discussion

10 Our work uses large-scale GWAS data to investigate the effect of genetically predicted education and
11 cardiometabolic risk factors on OA risk within the MR framework, and provides evidence supporting
12 protective effects of education and LDL-C and unfavourable effects of BMI and smoking. These
13 findings add insight into causal mechanisms underlying OA, its clustering with the risk factors of
14 cardiovascular disease, and disparities related to educational attainment.

15 Our results are consistent with previous MR analyses identifying a protective effect of genetically
16 predicted education and LDL-C, and a detrimental effect of genetically predicted BMI on OA risk (32-
17 34). However, our current study goes further to identify a novel association of genetically predicted
18 smoking with OA risk, and additionally quantify mediation of the effect of genetically predicted
19 education on OA risk through genetically predicted BMI and smoking. As higher education is
20 associated with lower LDL-C (4), this would not be consistent with LDL-C mediating the effect of
21 education on the risk of OA and therefore LDL-C was not considered in the mediation analysis. A
22 number of mechanisms have been proposed by which obesity and smoking might lead to increased
23 risk and severity of OA (35, 36). In contrast to our current findings, a meta-analysis of observational
24 studies has identified an inverse association between smoking and risk of knee OA (37). This

1 discrepancy may be attributable to limitations of conventional observational research for identifying
2 causal effects (38). Our current work also improves on a previous MR study exploring the causality of
3 smoking on OA risk, which only incorporated a single genetic variant to proxy smoking and found an
4 inverse association with risk of total joint replacement (39). This discrepancy may be explained by
5 our use of a greater number of instruments for smoking, to offer greater robustness against possible
6 violations of the MR modelling assumptions. Furthermore, our current study also considered OA
7 related to any joint, while the previous study only considered cases requiring hip or knee
8 replacement (39). As the pathophysiology of OA varies at different sites, this may also be
9 contributing to the observed differences in findings.

10 The findings of our study are relevant in both clinical and public health terms. Smoking and obesity
11 have widespread implications on human health that extend far beyond cardiovascular disease.
12 Smoking increases risk of chronic lung disease and many cancers, while obesity is a major
13 contributor towards risk of diabetes (40). Targeting of these risk factors therefore represents an
14 opportunity to simultaneously reduce risk of multiple distinct disease processes and thus ease the
15 burden of multi-morbidity on individuals and health systems alike (40). The identification of smoking
16 and obesity as downstream mediators of education supports that policies intended to increase
17 educational attainment should continue (4, 41). Educational attainment is known to be heritable,
18 and using variants robustly associated with the trait, we were able to explore associations with OA
19 risk. Previous work has suggested that it is the experience of being in education for longer
20 specifically, rather than related cognitive ability, that is likely deterministic of consequent health
21 outcomes (42).

22 Our results suggest that the protective effect of education on OA risk is mediated through smoking
23 and BMI. However, there was high uncertainty in the estimates, our data being consistent with the
24 mediated proportion being between 13% and 57%. For comparison, approximately half of the
25 protective effect of education on cardiovascular disease has previously been estimated to be

1 mediated together through blood pressure, obesity and smoking (3). Thus for OA more than
2 cardiovascular disease, education may be having a protective effect through pathways other than
3 downstream cardiometabolic mediators. Potential mechanisms underlying this may relate to
4 superior self-management and healthcare engagement practices afforded to those with greater
5 education (43, 44). Finally, our analyses also highlighted a potential protective effect of higher LDL-C
6 levels on OA risk. However, given the small magnitude of this, and particularly in relation to the
7 larger effect estimates seen for education, BMI and smoking (Figure 1), it is not clear that this is of
8 any clinical relevance.

9 Our study has limitations. Firstly, the MR approach uses the cumulative lifelong effect of genetic
10 variants and should not be extrapolated to presume the effect of a clinical intervention (45).
11 Secondly, the possibility of reverse causation that OA causes increased BMI or liability to smoking
12 cannot be completely ruled out. We did not examine the bidirectional associations because OA was
13 treated as a binary phenotype, and using such binary exposure is unlikely to capture the true causal
14 relationship in MR analysis (46). Thirdly, the OA and smoking genetic association estimates we use
15 were obtained using self-reported data, which may be subject to recall bias that could affect the MR
16 estimates generated (48). Fourthly, the UK Biobank cohort used to obtain many of the genetic
17 association estimates in this study represents a select group that may not be representative of more
18 general populations, and in particular non-European populations (49, 50). Fifthly, mediation analysis
19 crucially depends on the correct formulation of the causal relationships of the exposures *a priori*, as
20 mediation and confounding cannot be statistically distinguished(51). We assume adult BMI and
21 smoking mediate the effect of education, as supported by earlier literature (52, 53). Also,
22 interpreting mediation analysis results for a binary outcome is not straightforward due to the non-
23 collapsibility of the odds ratio, as the estimate for the mediated proportion may be biased (7).
24 Finally, we considered OA at any site in these analyses, and it is possible that the determinants of OA
25 vary across different anatomical locations (54).

1 In conclusion, this study uses genetic data in MR analyses to generate evidence supporting a
2 protective effect of education and detrimental effects of BMI and smoking on OA risk, with evidence
3 that the effect of education is mediated through BMI and smoking. These findings highlight
4 education, obesity and smoking as common mechanisms underlying clustering of OA with risk
5 factors of cardiovascular disease, which may represent clinical and public health targets for reducing
6 multi-morbidity and the burden of these common conditions.

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8 This research has been conducted using the UK Biobank Resource, and GIANT and GLGC genome-
9 wide association study summary data.

10 **Author contributions**

11 DG, and NS designed the study. DG, RM and VK performed statistical analyses. All authors
12 interpreted the results. DG, RM and NS drafted the manuscript. All authors edited the manuscript for
13 intellectual content. All authors take responsibility for the integrity of the study.

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2 The funding sources were not involved in study design, acquisition of data, analysis, interpretation or
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4 Conflicts of interest

5 DG is employed part-time by Novo Nordisk, outside the submitted work. NS has received
6 consultancy fees from Pfizer and Eli Lilly, but has no direct conflicts of interest relating to this
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8 References

- 9 1. OARSI. Osteoarthritis: A Serious Disease 2016 [cited 2020 February]. Available from:
10 <https://www.oarsi.org/education/oarsi-resources/oarsi-white-paper-oa-serious-disease>.
- 11 2. Wang H, Bai J, He B, Hu X, Liu D. Osteoarthritis and the risk of cardiovascular disease: a
12 meta-analysis of observational studies. *Sci Rep*. 2016;6:39672.
- 13 3. Carter AR, Gill D, Davies NM, Taylor AE, Tillmann T, Vaucher J, et al. Understanding the
14 consequences of education inequality on cardiovascular disease: mendelian randomisation study.
15 *BMJ*. 2019;365:l1855.
- 16 4. Tillmann T, Vaucher J, Okbay A, Pikhart H, Peasey A, Kubinova R, et al. Education and
17 coronary heart disease: mendelian randomisation study. *BMJ*. 2017;358:j3542.
- 18 5. Davey Smith G, Ebrahim S. Mendelian randomization: can genetic epidemiology contribute
19 to understanding environmental determinants of disease? *Int J Epidemiol*. 2003;32(1):1-22.
- 20 6. Burgess S, Thompson DJ, Rees JMB, Day FR, Perry JR, Ong KK. Dissecting Causal Pathways
21 Using Mendelian Randomization with Summarized Genetic Data: Application to Age at Menarche
22 and Risk of Breast Cancer. *Genetics*. 2017;207(2):481-7.

- 1 7. Carter AR, Sanderson E, Hammerton G, Richmond RC, Smith GD, Heron J, et al. Mendelian
2 randomisation for mediation analysis: current methods and challenges for implementation. *bioRxiv*.
3 2019:835819.
- 4 8. Bycroft C, Freeman C, Petkova D, Band G, Elliott LT, Sharp K, et al. The UK Biobank resource
5 with deep phenotyping and genomic data. *Nature*. 2018;562(7726):203-9.
- 6 9. Sudlow C, Gallacher J, Allen N, Beral V, Burton P, Danesh J, et al. UK Biobank: An Open
7 Access Resource for Identifying the Causes of a Wide Range of Complex Diseases of Middle and Old
8 Age. *PLOS Med*. 2015;12(3):e1001779.
- 9 10. Mbatchou J, Barnard L, Backman J, Marcketta A, Kosmicki JA, Ziyatdinov A, et al.
10 Computationally efficient whole genome regression for quantitative and binary traits. *bioRxiv*.
11 2020:2020.06.19.162354.
- 12 11. Pulit SL, Stoneman C, Morris AP, Wood AR, Glastonbury CA, Tyrrell J, et al. Meta-analysis of
13 genome-wide association studies for body fat distribution in 694 649 individuals of European
14 ancestry. *Hum Mol Genet*. 2019;28(1):166-74.
- 15 12. Willer CJ, Schmidt EM, Sengupta S, Peloso GM, Gustafsson S, Kanoni S, et al. Discovery and
16 refinement of loci associated with lipid levels. *Nat Genet*. 2013;45(11):1274-83.
- 17 13. Wootton RE, Richmond RC, Stuijzand BG, Lawn RB, Sallis HM, Taylor GMJ, et al. Evidence for
18 causal effects of lifetime smoking on risk for depression and schizophrenia: a Mendelian
19 randomisation study. *Psychol Med*. 2019:1-9.
- 20 14. Lee JJ, Wedow R, Okbay A, Kong E, Maghzian O, Zacher M, et al. Gene discovery and
21 polygenic prediction from a genome-wide association study of educational attainment in 1.1 million
22 individuals. *Nat Genet*. 2018;50(8):1112-21.
- 23 15. Okbay A, Beauchamp JP, Fontana MA, Lee JJ, Pers TH, Rietveld CA, et al. Genome-wide
24 association study identifies 74 loci associated with educational attainment. *Nature*.
25 2016;533(7604):539-42.

- 1 16. Burgess S, Davies NM, Thompson SG. Bias due to participant overlap in two-sample
2 Mendelian randomization. *Genet Epidemiol.* 2016;40(7):597-608.
- 3 17. Locke AE, Kahali B, Berndt SI, Justice AE, Pers TH, Day FR, et al. Genetic studies of body mass
4 index yield new insights for obesity biology. *Nature.* 2015;518(7538):197-206.
- 5 18. Hemani G, Zheng J, Elsworth B, Wade KH, Haberland V, Baird D, et al. The MR-Base platform
6 supports systematic causal inference across the human phenome. *eLife.* 2018;7.
- 7 19. Burgess S. Sample size and power calculations in Mendelian randomization with a single
8 instrumental variable and a binary outcome. *International Journal of Epidemiology.* 2014;43(3):922-
9 9.
- 10 20. Burgess S, Butterworth A, Thompson SG. Mendelian randomization analysis with multiple
11 genetic variants using summarized data. *Genet Epidemiol.* 2013;37(7):658-65.
- 12 21. Bowden J, Davey Smith G, Haycock PC, Burgess S. Consistent Estimation in Mendelian
13 Randomization with Some Invalid Instruments Using a Weighted Median Estimator. *Genet*
14 *Epidemiol.* 2016;40(4):304-14.
- 15 22. Burgess S, Foley CN, Allara E, Staley JR, Howson JMM. A robust and efficient method for
16 Mendelian randomization with hundreds of genetic variants. *Nat Commun.* 2020;11(1):376.
- 17 23. Bowden J, Davey Smith G, Burgess S. Mendelian randomization with invalid instruments:
18 effect estimation and bias detection through Egger regression. *Int J Epidemiol.* 2015;44(2):512-25.
- 19 24. Yavorska OO, Burgess S. MendelianRandomization: an R package for performing Mendelian
20 randomization analyses using summarized data. *Int J Epidemiol.* 2017;46(6):1734-9.
- 21 25. Burgess S, Thompson SG. Interpreting findings from Mendelian randomization using the MR-
22 Egger method. *Eur J Epidemiol.* 2017;32(5):377-89.
- 23 26. Slob EAW, Burgess S. A Comparison Of Robust Mendelian Randomization Methods Using
24 Summary Data. *bioRxiv.* 2019:577940.
- 25 27. Burgess S, Dudbridge F, Thompson SG. Re: "Multivariable Mendelian randomization: the use
26 of pleiotropic genetic variants to estimate causal effects". *Am J Epidemiol.* 2015;181(4):290-1.

- 1 28. Burgess S, Thompson SG. Multivariable Mendelian Randomization: The Use of Pleiotropic
2 Genetic Variants to Estimate Causal Effects. *American Journal of Epidemiology*. 2015;181(4):251-60.
- 3 29. Sanderson E. Multivariable Mendelian Randomization and Mediation. *Cold Spring Harbor
4 Perspectives in Medicine*. 2020.
- 5 30. Wasserstein RL, Schirm AL, Lazar NA. Moving to a World Beyond “ $p < 0.05$ ”. *The American
6 Statistician*. 2019;73(sup1):1-19.
- 7 31. Davey Smith G, Davies NM, Dimou N, Egger M, Gallo V, Golub R, et al. STROBE-MR:
8 Guidelines for strengthening the reporting of Mendelian randomization studies.
9 <https://doi.org/10.7287/peerj.preprints.27857v1>. *PeerJ Preprints*. 2019;7:e27857v1.
- 10 32. Zhu ZH, Zheng ZL, Zhang FT, Wu Y, Trzaskowski M, Maier R, et al. Causal associations
11 between risk factors and common diseases inferred from GWAS summary data. *Nat Commun*.
12 2018;9(1):224.
- 13 33. Hindy G, Akesson KE, Melander O, Aragam KG, Haas ME, Nilsson PM, et al. Cardiometabolic
14 Polygenic Risk Scores and Osteoarthritis Outcomes: A Mendelian Randomization Study Using Data
15 From the Malmo Diet and Cancer Study and the UK Biobank. *Arthritis Rheumatol*. 2019;71(6):925-
16 34.
- 17 34. Zengini E, Hatzikotoulas K, Tachmazidou I, Steinberg J, Hartwig FP, Southam L, et al.
18 Genome-wide analyses using UK Biobank data provide insights into the genetic architecture of
19 osteoarthritis. *Nat Genet*. 2018;50(4):549-58.
- 20 35. Amin S, Niu J, Guermazi A, Grigoryan M, Hunter DJ, Clancy M, et al. Cigarette smoking and
21 the risk for cartilage loss and knee pain in men with knee osteoarthritis. *Ann Rheum Dis*.
22 2007;66(1):18-22.
- 23 36. Powell A, Teichtahl AJ, Wluka AE, Cicuttini FM. Obesity: a preventable risk factor for large
24 joint osteoarthritis which may act through biomechanical factors. *Br J Sports Med*. 2005;39(1):4-5.
- 25 37. Kong L, Wang L, Meng F, Cao J, Shen Y. Association between smoking and risk of knee
26 osteoarthritis: a systematic review and meta-analysis. *Osteoarthr Cartilage*. 2017;25(6):809-16.

- 1 38. Felson DT, Zhang Y. Smoking and osteoarthritis: a review of the evidence and its
2 implications. *Osteoarthritis Cartilage*. 2015;23(3):331-3.
- 3 39. Johnsen MB, Vie GA, Winsvold BS, Bjorngaard JH, Asvold BO, Gabrielsen ME, et al. The causal
4 role of smoking on the risk of hip or knee replacement due to primary osteoarthritis: a Mendelian
5 randomisation analysis of the HUNT study. *Osteoarthritis Cartilage*. 2017;25(6):817-23.
- 6 40. Stewart ST, Cutler DM, Rosen AB. Forecasting the effects of obesity and smoking on U.S. life
7 expectancy. *N Engl J Med*. 2009;361(23):2252-60.
- 8 41. Di Chiara T, Scaglione A, Corrao S, Argano C, Pinto A, Scaglione R. Association between low
9 education and higher global cardiovascular risk. *J Clin Hypertens*. 2015;17(5):332-7.
- 10 42. Gill D, Efstathiadou A, Cawood K, Tzoulaki I, Dehghan A. Education protects against coronary
11 heart disease and stroke independently of cognitive function: evidence from Mendelian
12 randomization. *Int J Epidemiol*. 2019;48(5):1468-77.
- 13 43. Gustafsson K, Kvist J, Eriksson M, Dahlberg LE, Rolfson O. Socioeconomic status of patients in
14 a Swedish national self-management program for osteoarthritis compared with the general
15 population-a descriptive observational study. *BMC Musculoskelet Disord*. 2020;21(1):10.
- 16 44. Luong ML, Cleveland RJ, Nyrop KA, Callahan LF. Social determinants and osteoarthritis
17 outcomes. *Aging health*. 2012;8(4):413-37.
- 18 45. Gill D, Walker VM, Martin RM, Davies NM, Tzoulaki I. Comparison with randomized
19 controlled trials as a strategy for evaluating instruments in Mendelian randomization. *Int J*
20 *Epidemiol*. 2019.
- 21 46. Burgess S, Labrecque JA. Mendelian randomization with a binary exposure variable:
22 interpretation and presentation of causal estimates. *European Journal of Epidemiology*.
23 2018;33(10):947-52.
- 24 47. Palmer TM, Lawlor DA, Harbord RM, Sheehan NA, Tobias JH, Timpson NJ, et al. Using
25 multiple genetic variants as instrumental variables for modifiable risk factors. *Stat Methods Med*
26 *Res*. 2012;21(3):223-42.

- 1 48. Pulcu E. Self-report distortions of puffing topography in daily smokers. *J Health Psychol.*
2 2016;21(8):1644-54.
- 3 49. Batty GD, Gale CR, Kivimaki M, Deary IJ, Bell S. Comparison of risk factor associations in UK
4 Biobank against representative, general population based studies with conventional response rates:
5 prospective cohort study and individual participant meta-analysis. *BMJ.* 2020;368:m131.
- 6 50. Haworth S, Mitchell R, Corbin L, Wade KH, Dudding T, Budu-Aggrey A, et al. Apparent latent
7 structure within the UK Biobank sample has implications for epidemiological analysis. *Nat Commun.*
8 2019;10(1):333.
- 9 51. MacKinnon DP, Krull JL, Lockwood CM. Equivalence of the Mediation, Confounding and
10 Suppression Effect. *Prevention Science.* 2000;1(4):173-81.
- 11 52. Böckerman P, Viinikainen J, Pulkki-Råback L, Hakulinen C, Pitkänen N, Lehtimäki T, et al. Does
12 higher education protect against obesity? Evidence using Mendelian randomization. *Preventive*
13 *Medicine.* 2017;101:195-8.
- 14 53. Gage SH, Bowden J, Davey Smith G, Munafò MR. Investigating causality in associations
15 between education and smoking: a two-sample Mendelian randomization study. *International*
16 *Journal of Epidemiology.* 2018;47(4):1131-40.
- 17 54. Yucesoy B, Charles LE, Baker B, Burchfiel CM. Occupational and genetic risk factors for
18 osteoarthritis: a review. *Work.* 2015;50(2):261-73.

1 Table 1. Descriptive characteristics for the participants included in this study.

Variables	Osteoarthritis (including self report) N=80,646	Osteoarthritis (excluding self report) N=60,800	Controls N=328,250
Age, mean (SD), years	60.2 (6.7)	60.4 (6.8)	56.1 (8.1)
Sex, N (%)			
Male	33,084 (41.0)	26,073 (42.9)	154,802 (47.2)
Female	47,562 (59.0)	34,727 (57.1)	173,448 (52.8)
Never smoked, N (%)	39,876 (49.7)	29,697 (49.1)	182,523 (55.8)
Former smoker, N (%)	32,283 (40.2)	24,643 (40.7)	111,456 (34.1)
Current smoker, N (%)	8,086 (10.1)	6,149 (10.2)	33,234 (10.2)
BMI, mean (SD), kg/m ²	28.9 (5.3)	29.1 (5.3)	27.0 (4.6)
Incident cardiovascular events, N (%)	11,410 (14.1)	9,574 (15.7)	25,185 (7.7)
Diabetes diagnosed, N (%)	5,485 (6.8)	4,432 (7.3)	14,287 (4.4)
Systolic blood pressure, mmHg (SD)	140.6 (18.4)	140.6 (18.4)	137.8 (18.7)
Diastolic blood pressure, mmHg (SD)	82.5 (9.9)	82.6 (9.9)	82.3 (10.2)

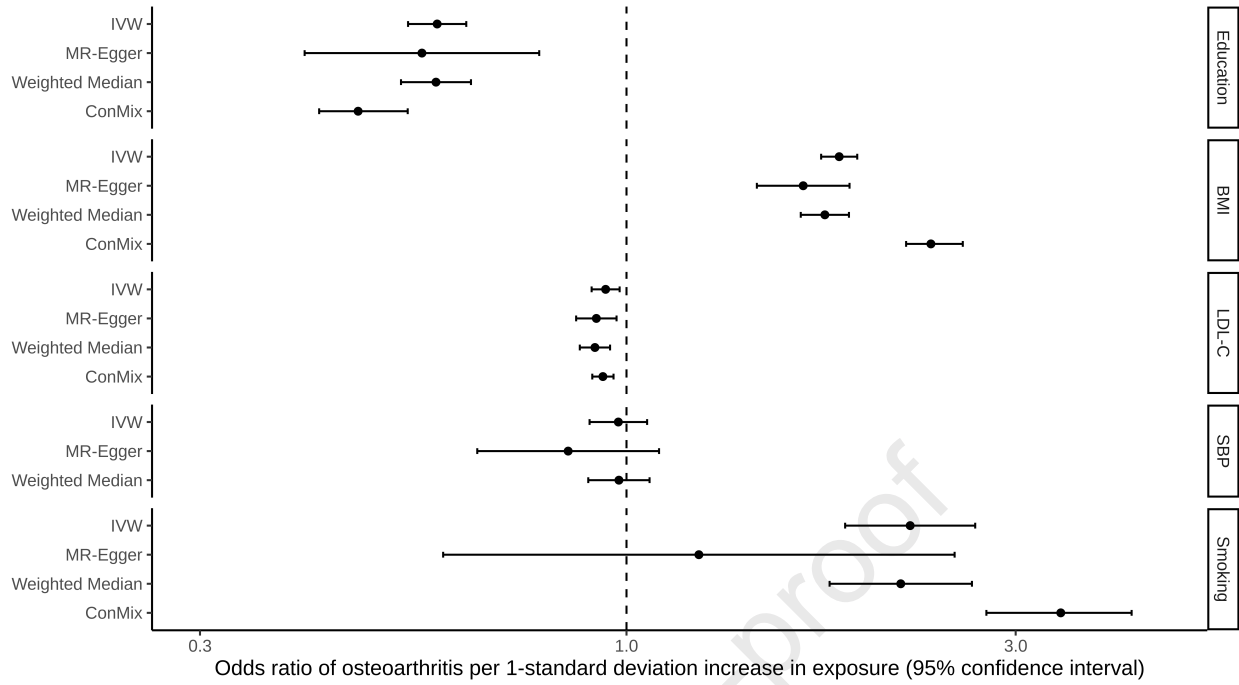
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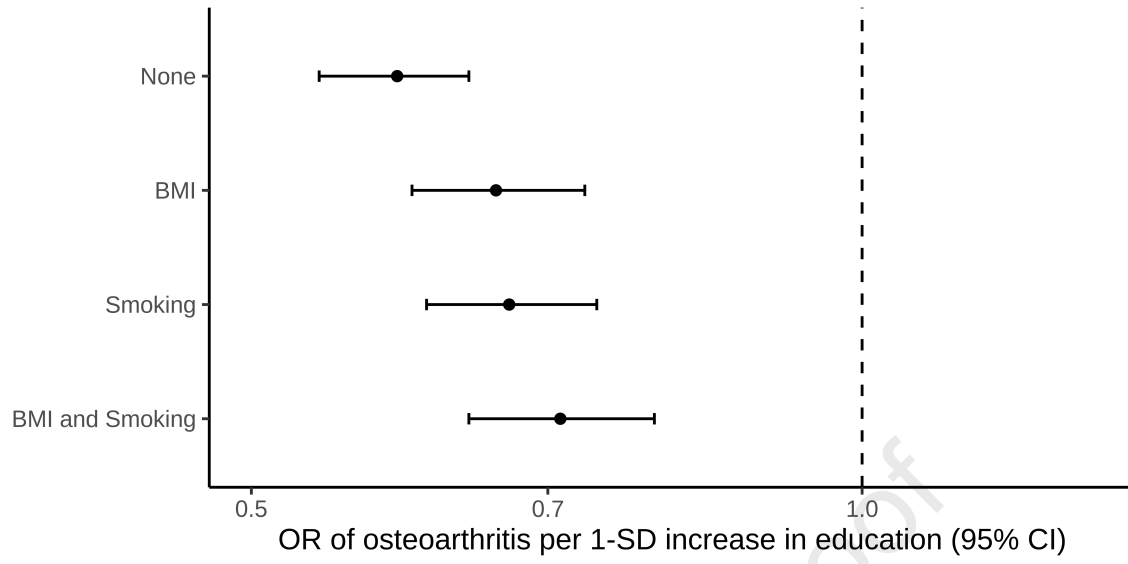
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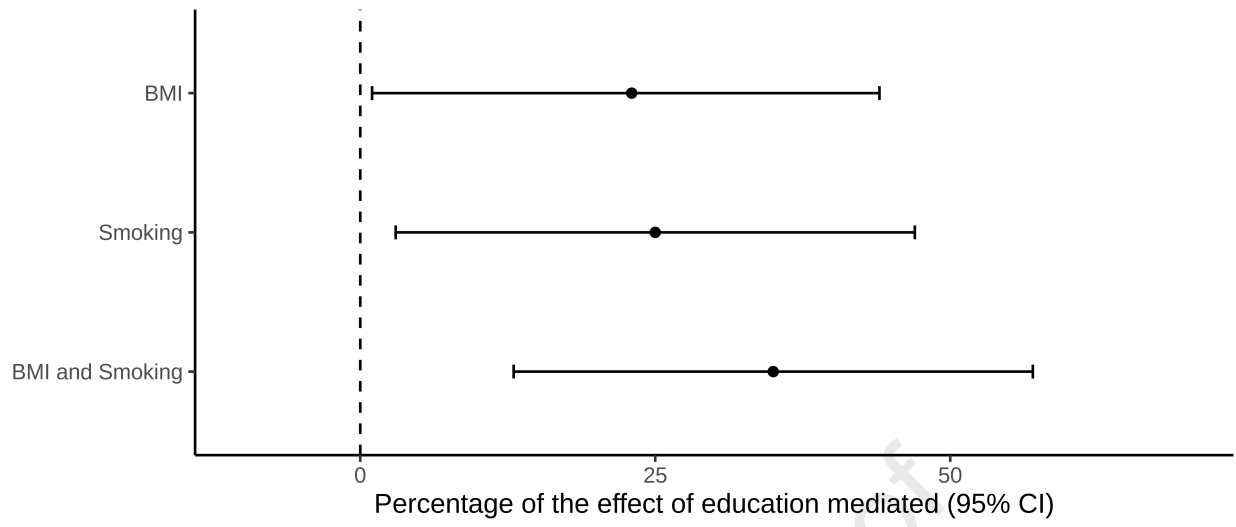
2 **Figure 1.** Effects of genetically predicted education, body-mass index (BMI), low-density lipoprotein
3 cholesterol (LDL-C), systolic blood pressure (SBP) and lifetime smoking respectively on risk of
4 osteoarthritis. Inverse-variance weighted (IVW), contamination mixture (Con-Mix), Egger and
5 weighted median represent different Mendelian randomization models. Confidence intervals could
6 not be generated for the Con-Mix analysis considering SBP, and hence this result is not presented.

7 **Figure 2.** The effect of genetically predicted education on osteoarthritis risk after adjusting for
8 genetically predicted body-mass index and lifetime smoking, either separately or in the same model.
9 The y-axis details the adjustment made. CI: confidence interval; OR: odds ratio; SD: standard
10 deviation.

11 **Figure 3.** The percentage of the effect of genetically predicted education on osteoarthritis risk that is
12 mediated through genetically predicted body-mass index (BMI) and lifetime smoking, separately and
13 when considered together in the same model. The y-axis details the mediating pathway considered.
14 CI: confidence interval.







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