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- 1 Development and evaluation of a gentamicin pharmacokinetic model that facilitates
- 2 opportunistic gentamicin therapeutic drug monitoring in neonates and infants.
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Abstract

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44 45 Trough gentamicin therapeutic drug monitoring (TDM) is time-consuming, disruptive to neonatal clinical care and a patient safety issue. Bayesian models could allow TDM to be performed opportunistically at the time of routine blood tests. This study aimed to develop and prospectively evaluate a new gentamicin model and a novel Bayesian computer tool (neoGent) for TDM use in neonatal intensive care. We also evaluated model performance for predicting peak concentrations and AUC(0-t). A pharmacokinetic meta-analysis was performed on pooled data from three studies (1325 concentrations from 205 patients). A 3-compartment model was used with covariates being: allometric weight scaling, postmenstrual and postnatal age, and serum creatinine. Final parameter estimates (standard error) were: clearance: 6.2 (0.3) L/h/70kg; central volume (V) 26.5 (0.6) L/70kg; inter-compartmental disposition: Q=2.2 (0.3) L/h/70kg, V2=21.2 (1.5) L/70kg, Q2=0.3 (0.05) L/h/70kg, V3=148 (52.0) L/70kg. The model's ability to predict trough concentrations from an opportunistic sample was evaluated in a prospective observational cohort study that included data from 163 patients with 483 concentrations collected in five hospitals. Unbiased trough predictions were obtained: median (95% confidence interval (CI)) prediction error was 0.0004 (-1.07, 0.84) mg/L. Results also showed peaks and AUC(0-t) could be predicted (from one randomly selected sample) with little bias but relative imprecision with median (95% CI) prediction error being 0.16 (-4.76, 5.01) mg/L and 10.8 (-24.9, 62.2) mg h/L, respectively. NeoGent was implemented in R/NONMEM, and in the freely available TDMx software.

Introduction

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The aminoglycoside antibiotic gentamicin is the most commonly used antimicrobial on neonatal 47 48 units(1, 2) and is effective against Gram negative bacteria. Gentamicin use is limited by its narrow 49 therapeutic index and risk of toxicity, specifically nephro- and ototoxicity(3). It is not metabolized in the liver(4) and is almost entirely eliminated by the kidneys; clearance therefore depends on renal 50 function. During the first two weeks of life, renal and intra-renal blood flow increase rapidly, causing 51 52 a steep rise in glomerular filtration rate (GFR)(5, 6). Therapeutic drug monitoring (TDM) is required to ensure maximal efficacy and especially minimal 53 toxicity, particularly in the neonatal population where variability in pharmacokinetic (PK) parameters 54 55 is large. Dose individualization approaches focus on toxicity(7, 8) and include single-level methods 56 and nomograms(9, 10), area under the curve (AUC) methods(11), and Bayesian methods(12). The use 57 of nomograms is limited as they cannot readily incorporate covariates affecting PK parameters. AUC methods use a simplified 1-compartment PK model and require at least two gentamicin 58 59 measurements, which is not appropriate in neonates with limited blood volumes. These drawbacks 60 make Bayesian approaches the most attractive for newborn infants. Deriving a Bayesian prior for TDM requires a non-linear mixed-effect PK model, and several such 61 62 studies of neonatal gentamicin have been published(13-24). However, these studies are limited by their heterogeneity and use of sparse data (often identifying only a 1-compartment model when 63 gentamicin follows multi-compartment kinetics(25, 26)) and fail to account for age-related differences 64 65 in creatinine during the immediate newborn period. Although gentamicin is not a new drug, its dosing and monitoring is still a current issue as identified in the UK National Patient Safety alert 66 67 (http://www.nrls.npsa.nhs.uk/alerts/?entryid45=66271) and a recent publication by Valitalo et al(27), 68 who used simulations to define dosing guidelines. 69 We aimed to investigate whether opportunistic sampling can predict trough gentamicin concentrations 70 so that standard TDM could be performed from a blood sample taken for other purposes (e.g. routine blood gases). As a secondary aim, we evaluated the model's ability to predict peak gentamicin 71 72 concentrations and AUC(0-t) using one randomly selected sample.

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74	Study	popula	ation

75 This study used two datasets: a model-building dataset and a prospectively collected evaluation

76 dataset.

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To collect data for model development, the electronic bibliographic database PubMed was searched in 77

January 2015 without time limitations. The search strategy included: (neonat* OR newborn*) AND

(gentamicin) AND (pharmacokinetic* OR PK); gentamicin samples had to be prospectively collected

80 and covariates (weight, gestational age (GA), postnatal age (PNA), serum creatinine measurements),

also had to be reported. Additionally, we also searched the reference lists in identified papers. The

authors of the publications that met the inclusion criteria (n=8) (11, 15, 21, 22, 28-31) were then 82

83 invited to contribute their data.

84 Data for the evaluation of the PK model were collected as a prospective observational cohort study

from five UK hospitals (St George's University Hospitals NHS Foundation Trust, Liverpool Women's 85

86 NHS Foundation Trust, Oxford University Hospitals, Portsmouth Hospitals NHS Trust and Coventry

87 & Warwickshire University Hospitals NHS Trust) from July 2012 to November 2013. Infants were

eligible for inclusion if the following criteria were met: more than 36 hours gentamicin therapy 88

89 anticipated, postnatal age of less than 90 days, not receiving extracorporeal membrane oxygenation,

peritoneal dialysis or hemofiltration, and expected to survive the study period (as judged by the 90

91 clinical team). Each patient provided a minimum of two gentamicin concentrations - a trough sample

92 from routine TDM (i.e. a pre-dose sample taken before a non-initial dose) and an additional study

93 sample (taken opportunistically during a course of gentamicin when the infant required blood

94 sampling for clinical care). These samples will be referred to as routine (trough) and (opportunistic)

95 study samples in this manuscript. Exact times of gentamicin dosing and sampling were recorded,

96 along with the patient's weight, age and serum creatinine (Table 1). Written informed consent was

obtained from parents and the study was approved by the London Central Ethics committee (reference

12/LO/0455).

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Gentamicin dosing and sampling procedure in the prospective evaluation dataset

101 Gentamicin treatment was initiated at the discretion of the clinical team for possible infection and 102 dosed and monitored using trough concentrations according to the standard practice at each hospital. 103 Gentamicin was administered as a slow (<2 min) bolus via intravenous cannula, percutaneous long 104 line, or umbilical venous catheter. 105 106

Bioanalytical techniques

107 An enzyme immunoassay (EMIT, Syva)(15), a fluorescence polarization immunoassay (TDx, Abbot)(15, 21), and high performance liquid chromatography coupled to tandem mass spectrometry 108 109 (UHPLC-MS/MS) (32) were used to determine gentamicin concentration in the model-building 110 dataset; and the Jaffe reaction (33) was used to determine serum creatinine concentrations. In the 111 prospective evaluation dataset, gentamicin serum concentrations were analyzed using immunoassay 112 techniques (Table S1); and creatinine concentrations were determined by either a Jaffe-based or an 113 enzymatic method (137 neonates and 26 neonates, respectively).

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Pharmacokinetic analysis

The observed concentration-time data from only the model-building studies were pooled and 116 117 simultaneously analyzed with non-linear mixed-effects software NONMEM version 7.3(34). The first 118 order conditional estimation method with interaction was used.

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- 120 Basic model
- One-, 2-, and 3-compartment structural models were considered when defining the basic structural 121
- 122 population PK model. The inter-individual variability (IIV) was assumed to follow a log-normal
- 123 distribution and tested on all parameters. An additive, a proportional, and a combination of both
- 124 (Equation 1) residual error models were tested.
- $y_{ij} = f(t_{ij}; \phi_i) + f(t_{ij}; \phi_i) \cdot \varepsilon_{ij(proportional)} + \varepsilon_{ij(additive)},$ (Equation 1) 125
- 126 where y_{ij} is an observed gentamic concentration at time t_{ij} , f is the function that represents the
- gentamic model, ϕ_i is a vector of parameters, ε_{ij} is a residual error term. 127

Inter-occasion variability (IOV) was also assumed to be log-normally distributed and it was tested for all parameters with an occasion defined as a single dosing interval.

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131 Covariate model

- Allometric scaling was used a priori to standardize all PK parameters to 70 kg (35), and a maturation 132 function, describing the maturation of the GFR with postmenstrual age (PMA) (Equation 2) with fixed 133 134 parameters from a previous study (5), was used to scale clearance. Allometric exponents were fixed to 0.632 for central clearance and 0.75 for inter-compartmental clearances. Different exponents were 135 136 used because these values were shown best for describing the maturation of renal elimination(5) and 137 tissue blood flows(36), respectively. Allometric exponents for volumes of distribution were fixed to 1. 138 The combination of allometric weight scaling and sigmoidal maturation function was suggested as a 139 standard method for scaling clearance in the pediatric population in a recent comparison of different approaches(37). 140
- $maturation function = \frac{PMA^{Hill}}{PMA_{50}^{Hill} + PMA^{Hill}},$ 141 (Equation 2)
- where Hill is the sigmoidicity coefficient and PMA₅₀ is PMA when maturation of GFR reaches 50% 142
- 143 of adult values.
- 144 As it is known that PNA and serum creatinine are important indicators of gentamicin clearance and 145 also based on the posthoc estimates of etas versus covariates plots, they were tested on clearance.
- 146 These time-varying covariates were considered to significantly improve the fit and therefore included
- in the model if the difference in objective function value (ΔOFV) after their inclusion was >3.84 147
- (p<0.05). Additionally, linear extrapolations between observations were made. To account for 148
- 149 endogenous creatinine, maternal creatinine and also the change in renal function with age, a typical
- 150 value of serum creatinine (TSCr) for a specific PMA was determined using data from Cuzzolin et
- al(38) for preterm (GA<37 weeks) newborns and Rudd et al(39) for term newborns. A linear decline 151
- in TSCr with increasing PMA was found according to Equation 3: 152
- $TSCr = -2.849 \cdot PMA (weeks) + 166.48.$ (Equation 3) 153

154 A possible influence of serum creatinine on clearance was tested according to the following Equation

4, where measured serum creatinine (MSCr) was standardized by TSCr for PMA and departures from

156 it estimated as follows:

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$$\left(\frac{MSCr}{TSCr}\right)^{\theta}$$
. (Equation 4)

158 The effect of PNA was investigated with a logistic function (Equation 5) to account for the rapid

changes in gentamicin clearance in the first hours of life. The first day of life was defined as day 1. 159

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$$postnatal\ age\ function = \frac{PNA}{PNA_{50} + PNA},$$
 (Equation 5)

where PNA₅₀ is the PNA when clearance has reached 50% of typical adult's clearance. 161

162 After the forward selection (ΔOFV>3.84) of all covariates (full model), backward elimination was

163 performed, with a p-value retention cut-off of 0.001 (Δ OFV<10.83).

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165 Evaluation

166 Internal model evaluation

Basic goodness-of-fit plots for observations versus population and individual predictions, conditional 167

168 weighted residuals versus population predictions and versus time after dose were produced using

statistical software R version 3.1.0 (R Core Team (2014). R: A language and environment for 169

170 statistical computing. R Foundation for Statistical Computing, Vienna, Austria. Available from:

171 http://www.R-project.org/) and visually examined. The assumptions of normality and homogeneity of

172 the residuals errors were investigated by inspecting a histogram and a qq-plot.

173 Standard errors from NONMEM covariance step and non-parametric bootstrap analysis with 1,000

174 replicates were used to determine the precision of the final PK parameter estimates.

Additionally, we simulated 1,000 datasets using parameter estimates from the final model, and plotted 175

95% confidence intervals (CI) around the 2.5th, 50th, and 97.5th prediction percentiles of the simulated 176

data. Then, the observations were overlaid on the plot, also called the visual predictive check (VPC). 177

Perl-speaks-NONMEM (PsN) software(40) was used for the bootstrap analysis and to produce the 178

179 VPC, which was visualized using R-package Xpose4(41).

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181 External model evaluation

The prospective evaluation dataset was used to evaluate the predictive performance of the model. No 182

183 additional fitting was done, and the diagnostic plots and the VPC were generated as described above.

184 Bayesian model-predicted trough concentrations were computed using the model as a prior and

information from only the opportunistic study samples. These predictions were compared with the 185

observed trough concentrations by calculating the prediction error (PE) (42), and also the mean PE 186

187 (MPE) (i.e. a measure of bias), and root-mean-square error (RMSE), a measure of precision(43)

188 (Equations 6).

PE = observed - predicted189

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$$MPE = \frac{1}{N} \cdot \sum_{i=1}^{N} \cdot PE_i$$
 (Equations 6)

$$RMSE = \sqrt{\frac{1}{N} \cdot \sum_{i=1}^{N} \cdot PE_i^2}$$

Also, we counted the number of "correct" predictions that were below or above the currently 191

192 recommended gentamicin trough concentration thresholds of 1 mg/L or 2 mg/L (the National Institute

for Health and Care Excellence (NICE) (http://www.nice.org.uk/guidance/CG149/chapter/1-193

194 Guidance#therapeutic-drug-monitoring-for-gentamicin) and British National Formulary for Children

195 (BNFc) (http://www.evidence.nhs.uk/formulary/bnfc/current/5-infections/51-antibacterial-drugs/514-

196 aminoglycosides/gentamicin)).

Further analysis of paired samples (that is both study and routine samples taken in the same dosing 197

198 interval) was undertaken for the following scenarios: study samples ≥1, ≥2, and ≥3 mg/L, compared

199 with only unpaired samples.

201 Cross-validation

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202 The subset with the study sample above 3 mg/L provided the most important comparison, since in this

203 case the study sample was still above the pre-specified trough threshold. As there were only 18 pairs

204 with opportunistic study concentration ≥3 mg/L in the evaluation dataset, these pairs were merged

205 with paired samples of the same characteristics from the model-building dataset. The pooled dataset

was then randomly split into five subsets, and cross-validation was performed; meaning that in each subset 20% of the pairs were randomly removed and the model was re-estimated. The re-estimated model was then used as a prior to predict the troughs, and compared to the observed trough concentrations as previously described. Whether the model is able to predict peak concentrations from one randomly selected non-peak sample was tested similarly as described above, using paired samples from both the model-building and the evaluation dataset, and performing cross-validations. Additionally, as a possible pharmacokinetic-pharmacodynamic target for aminoglycosides can also be AUC(0-24)/MIC (44), the model was also evaluated on how it predicts AUC(0-t). Only a subset of the data where five or more samples were collected after the same dose was used for defining AUC(0-t), and the model-predicted versus observed (non-compartmental) AUC(0-t) was compared. Comparison with other models To compare our mechanistic model which scales for size, age and expected renal function with previously published models using empirical covariate analysis, predictions for the measured trough from the routine opportunistic samples in our prospective dataset were generated.

neoGent software

The model was implemented using R and NONMEM (see Supplementary material). It works by reading an individual's data into R, then Bayesian estimates generated in NONMEM are used to predict outcomes of interest (e.g. the time when the concentration falls below 2 mg/L).

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229 Results

230 **Patients**

231 Out of eight contacted authors identified in the literature search we obtained two large neonatal 232 gentamicin datasets (15, 21). We received no response from four authors (11, 28-30); and although an 233 initial response was received from two authors (22, 31) no data were actually shared. Additionally, we obtained some previously unpublished data taken during a PK study of ampicillin and penicillin (32). 234 235 The data were pooled and comprised 1325 gentamicin concentrations from 205 neonates (Table 1). 236 This dataset was used to derive the model. 237 For the model evaluation, gentamicin serum concentrations were prospectively collected from a total 238 of 194 neonates. Of the enrolled patients, 163 were included in the PK analysis (Table 1). Reasons for 239 exclusion (31 patients) included inexact sampling times, insufficient samples, or the gentamicin

240 opportunistic study concentration being below the limit of quantification (n=12). The final evaluation dataset comprised 483 gentamicin serum measurements, with 229 study and 254 routinely taken 241

242 trough concentrations. Median (range) time after dose was 13.3 (0.08-53.3) h and 31.1 (8.0-79.7) h for

study and routine concentrations, respectively. Patients were on treatment for up to 20 days.

245 Pharmacokinetic analysis

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Initially, a 2-compartment model provided a better fit to the data (ΔOFV=7.4 with a 3-compartment model) and was therefore chosen as the basic structural model. But, after the addition of the fixed allometric and renal function parameters, covariates and IOV, a 3-compartment model described the data better (47-unit drop in OFV). The IIV was described with an exponential error structure, and the best residual error model was a combination of a proportional and additive error.

251 Postnatal age and standardized serum creatinine had a significant effect on clearance (ΔOFV=134.1 252 and $\Delta OFV=17.2$, respectively) and were thus included in the final model. Backward elimination 253 (p=0.001) confirmed that these covariates remained significant with the 3-compartment model. The

final gentamicin population PK model is summarized with Equations 7. 254

$$255 \qquad CL = \theta_{CL} \cdot \left(\frac{WT}{70}\right)^{0.632} \cdot \frac{PMA^{3.33}}{55.4^{3.33} + PMA^{3.33}} \cdot \left(\frac{MSCr}{TSCr}\right)^{\theta_{SCr}} \cdot \frac{PNA}{\theta_{P_{50}} + PNA} \cdot e^{(\eta_{CL} + \kappa_{CL})},$$

$$\begin{split} V &= \; \theta_V \cdot \left(\frac{wT}{70}\right) \cdot e^{\eta_V}, \\ Q &= \; \theta_Q \cdot \left(\frac{wT}{70}\right)^{0.75} \cdot e^{\eta_Q}, \end{split}$$
256 (Equations 7)

- 257
- 258 where CL is gentamicin clearance, V is gentamicin volume of distribution, O is inter-compartmental 259
- 260 gentamic clearance, WT is body weight in kilograms, η is IIV, κ is IOV.
- There was only a small improvement in fit (ΔOFV=7.6) when the model was parameterized for time-261
- 262 varying covariates (linear extrapolation between observed covariate values), but as this model is more
- 263 biologically plausible, it was chosen as the final model.
- The OFV reduced from 2305.0 to 1217.5 between the basic and the final model. The inclusion of the 264
- covariates resulted in a reduction of the IIV on PK parameters: with the basic model the IIV on CL 265
- 266 and V was 71.1% and 62.5%, respectively, and with the final model, 41.8% and 33.5%, respectively.
- The final PK parameter estimates with uncertainty are reported in Table 2. 267
- 269 **Evaluation**

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- 270 Internal model evaluation
- 271 Figure 1 shows plots assessing goodness-of-fit by comparing observations and predictions. A VPC of
- 272 the final model is shown in Figure 2.
- 274 External model evaluation
- The basic diagnostic plots are presented in Figure 1, and the VPC performed using the evaluation 275
- dataset and the final parameters from the PK model without additional fitting in Figure 2. 276
- 277 Table 3 shows the number of correct predictions (for five different datasets from the evaluation data
- and pooled results from the cross-validation) for gentamicin trough thresholds of 1 and 2 mg/L 278
- 279 together with prediction errors. In the total dataset, containing both paired and unpaired samples, the
- median (95% CI) PE was 0.0004 (-1.1, 0.8) mg/L. The MPEs when predicting trough and peak 280
- 281 concentrations (using cross-validations) were 0.03 and 0.19 mg/L; and the RMSE 1.28 and 2.55 mg/L,
- 282 respectively (Table 3). When AUC(0-t) prediction (from one random sample) was evaluated, MPE
- was 14.5 mg h/L, and RMSE 30.2 mg h/L. 283

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Figure 3 shows the median and the range of PE for this model and previously published gentamicin 284 285 population PK models. 286 NeoGent 287 288 Figure S1 shows an example of output from neoGent. 289

Discussion

A PK model for gentamicin in neonates was developed and evaluated with prospectively collected data. Through its use of mechanistic covariates the model gave unbiased predictions of trough concentration from an opportunistic sample. Using this model, concentrations from samples taken at any time can be used to generate informative TDM, potentially eliminating the need for specifically timed trough gentamicin samples and the safety concerns and inconvenience associated with them. An exploratory analysis to evaluate whether such an approach could be used for predicting individual peak concentration and AUC(0-t) showed that while predictions were unbiased, they were relatively imprecise (Table 3).

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> The small median PE (0.0004 mg/L) for trough concentrations suggests that the model implemented in neoGent performs well, although some outliers were not captured (range: -2.4 - 1.6 mg/L). The median prediction errors were in most cases negative (Table 3), indicating that the model slightly over-predicts the trough concentrations (i.e. predicts them to be higher than they are), which might be (from a safety perspective) preferable to under-predicting. Cross-validations confirmed that samples do not need to be taken at a specific time when using this model for TDM, as predictions of trough concentrations (using an opportunistic sample) were unbiased, with median PE of -0.04 mg/L (Table 3). Although we did not test the effect of the sampling time on model predictions; the samples were collected from a wide range of times (0.1-53.3 h after the dose), as they would be in routine hospital tests.

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Comparison of the developed model with the existing published models showed that the predicted trough concentrations were the least biased (i.e. the median prediction error was the smallest) when our model was used (Figure 3). However, due to unavailability of some covariates in our dataset, three models were used without all of the covariates (APGAR score(15, 19), sepsis(19), co-medication with dopamine(23)) included, which could explain their worse predictive performance.

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The rich data in our model-building dataset (6.5 samples per patient) supported a 3-compartment model, where the final estimates for the third compartment were: inter-compartmental clearance 0.3 L/h/70kg and peripheral volume of distribution of 148 L/70kg. Additionally, the terminal half-life for a typical subject from the prospective evaluation dataset (weight 2.0 kg, PMA 34.9 weeks, PNA 6 days, MSCr 47.0 µmol/L, TSCr 66.4 µmol/L) was 189.7 hours. This could indicate uptake of gentamicin into the renal cortex, and slow excretion from it (45); and is in agreement with previously found evidence of deep tissue accumulation of gentamicin (26, 46).

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> Unfortunately many authors were unwilling or unable to share their data and we only managed to obtain data from two (15, 21) out of eight identified studies for our model building dataset. We did obtain one further subsequent dataset where assays from another pharmacokinetic study in neonates also receiving gentamicin were used (32). Due to differences in model structure and parameterization, it was not possible to extract relevant information for model building from the published reports. However, in part because data from Nielsen et al(21) was of such high quality with multiple samples per patient, our final model described both model building and the evaluation datasets well, as shown in Figures 1 and 2. The histogram and the qq-plot of the conditional weighted residuals (data not shown) confirmed that they follow a normal distribution. The final estimates for clearance (CL) and volume of distribution (V) were (mean (standard error)) 6.21 (0.30) L/h/70 kg and 26.5 (1.11) L/70kg, respectively (Table 2). The values of the PK parameters for a typical infant from the model-building dataset (weight 2.12 kg, PMA 33.0 weeks, PNA 5.4 days, MSCr 78 µmol/L, TSCr 71.4 µmol/L) were 0.077 L/h and 0.80 L (and 0.10 L/h and 0.78 L for a neonate from the evaluation dataset) for CL and V, respectively. These values are in agreement with estimates for clearance from previous neonatal studies of gentamicin pharmacokinetics(13, 14, 18, 22-24). The reported value for CL from Nielsen et al(21) may appear to be lower (0.026 L/h), but when our median demographic values were used in their model, the CL became similar to our estimates (0.095 L/h). The final estimate for volume of distribution is consistent with the estimate from Fuchs et al(23) and Botha et al(24), but it is not in accordance with what was found by Garcia et al (20) (0.252 L). The probable reason for this is a

different studied population, as when the median weight from our dataset was used in their model, the resulting V was 0.968 L, in agreement with our estimate.

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We did not attempt to estimate the allometric power exponents and constants of the maturation function as the PMA in the studied neonates (23.3-43.8 weeks) was insufficient to capture the age when maturation is complete (PMA₅₀=55.4 weeks(5)); instead, these constants were fixed to the values from another study in which the main focus was renal maturation(5). This type of scaling was used to improve the model usefulness by allowing it to be extrapolated to different subpopulations (for example, neonates with a different weight, or PMA). In addition to changes in clearance due to long-term maturation that extends throughout gestation and into the first two years of life, we attempted to capture the short-term changes in clearance that occur after birth regardless of gestational age. A benefit of fixing the long-term maturation based on known relationships between PMA and renal function was that this short-term maturation was apparent with our estimate of PNA₅₀ of 40.8 hours, indicating that clearance rapidly increases over the first few days of life. In the first day of life the clearance was at 37% of the value for a typical adult, and it reached 95% by the end of the first month of age.

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The typical serum creatinine (used in the model) was determined using SCr concentrations, determined by the Jaffe assay, because the same method was used to determine SCr in the modelbuilding dataset. But to determine SCr in the evaluation dataset, assays, based on both the Jaffe and the enzymatic methods, were used. However, the goodness-of-fit to the evaluation dataset and the predictive performance of the model were good, therefore no correction factor was included. Also, the enzymatic assay was only used in 16% of patients. Due to the range of the data that was used to determine typical-for-PMA SCr the model can be used for a neonate with PMA <44 weeks or a term neonate of <4weeks of age. The power exponent on the creatinine function was estimated to be -0.13, meaning that if observed SCr and typical SCr were 70 μmol/L and 60 μmol/L, respectively, clearance would be 2% lower.

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Large η-shrinkage indicates that the data do not contain enough information to make a reliable individual estimation. And whilst the shrinkage was large on the peripheral volumes of distribution (V2 and V3), it was relatively small on clearance (6.9%) (Table 2), which is important for making predictions of trough gentamicin concentrations and AUC(0-t). The η-shrinkage was also relatively small (15%) on the central volume of distribution (Table 2).

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Although the main aim was to evaluate whether the model can predict trough concentrations, the ability of the model to predict peak gentamicin concentration (from a randomly-selected non-peak sample) was also examined. Cross-validations showed that the median prediction error (95% CI) when predicting peaks was 0.16 (-4.76, 5.01) mg/L, indicating unbiased, but not very precise predictions. This is perhaps not surprising, given that concentrations collected at a median time after dose of 19.3 hours were used to predict concentrations at median 1h post dose. The prediction of AUC(0-t) (also from one sample) was similarly unbiased (median prediction error 10.8 mg h/L), but imprecise (95% CI: -24.9, 62.2 mg h/L) (Table 3). However, normalized RMSEs (by the range of observed data) for peak and AUC(0-t) prediction were 7.0% and 17.6%, respectively; indicating that considering the range of possible values, the precision is perhaps more acceptable. Target AUC(0-24) or peak values have not been defined in neonates, and slow clearance and a narrow therapeutic index mean that adjusting doses to target efficacy in this population may not be realistic. However, our model does now give unbiased predictions of both metrics from an opportunistically collected single sample, which should prove useful in future clinical research to define efficacy targets in this age group. At present, due to their imprecision, these predictions (for peak concentration and AUC(0-t)) should currently only be used for research purposes, and not for dose adjustment.

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Conclusion

A new gentamicin model has been developed and evaluated with prospectively collected data. We used mechanistic covariate parameterization informed by principles of allometric size scaling, known scaling of glomerular filtration maturation, and standardization for age-expected creatinine. This "biological prior" information gave a model with better predictive performance on prospectively

collected external data than any previously published gentamicin model. Using this we developed a software tool neoGent (see Supplementary material for provisional stand-alone version, and implemented in the web TDM application TDMx (http://www.tdmx.eu/) (47)), which can be used to predict when the trough concentration will fall below 2 mg/L and so guide the dosing interval. Furthermore, peak concentration or AUC(0-24) from any post-dose sample can also be predicted with little bias.

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Transparency declarations

426 None to declare.

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Tables and figures 547

548

549 Table 1: A summary of demographics and dosing

	Model-building dataset	Evaluation dataset
n	205	163
weight (g) ^a	2.12 (0.53-5.05)	2.03 (0.48-5.05)
gestational age (weeks) ^a	34.0 (23.3-42.1)	34.3 (23.9-42.3)
postnatal age (days) ^a	5.4 (1-66)	6 (1-78)
postmenstrual age (weeks) ^a	33.0 (23.3-43.8)	34.9 (24-43.3)
females (%)	89 (43%)	68 (41.7%)
gentamicin samples per patient b	6.5	3.0
gentamicin concentration (mg/L) ^a	3.4 (0.3-37.6)	1.0 (0.1-13.2)
time after the dose (h) ^a	8.0 (0.02-54.1)	23.5 (0.08-79.7)
occasion ^a	2 (1-22)	2 (1-7)

Weight and gestational age are values at treatment initiation, the rest are values at time of gentamicin 550

sampling/dosing; an occasion was defined as a dose with subsequent gentamicin samples taken; day 551

552 of birth was defined as day 1; amedian (range); mean

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Table 2: Final parameter estimates from NONMEM output file and from the bootstrap analysis 554

	Parameters from the final model				Bootstrap analysis		
	mean	SE	%CV	η-shrinkage	median	2.5%ile	97.5%ile
CL (L/h/70kg)	6.21	0.30	-	-	6.14	5.47	6.75
θ_SCr	-0.13	0.055	-	-	-0.13	-0.25	-0.03
PNA ₅₀ (days)	1.70	0.30	-	-	1.68	1.15	2.30
V (L/70kg)	26.5	1.11	-	-	26.3	23.6	28.4
Q (L/h/70kg)	2.15	0.32	-	-	2.19	1.68	3.25
V2 (L/70kg)	21.2	1.50	-	-	20.9	17.9	24.2
Q2 (L/h/70kg)	0.27	0.047	-	-	0.28	0.19	0.38
V3 (L/70kg)	148	52.0	-	-	152	65.2	534
IIV on CL	0.175	0.038	41.8	6.9	0.170	0.104	0.254
IIV on V	0.112	0.032	33.5	15.2	0.113	0.057	0.190
covariance CL-V	0.116	0.030	-	-	0.115	0.060	0.184
IIV on V2	0.132	0.060	36.3	57.8	0.117	0.023	0.281
IIV on V3	0.177	0.216	42.1	85.0	0.114	0.00002	4.18
inter-occasion variability	0.014	0.007	11.8	-	0.013	0.001	0.029
residual error (proportional)	0.036	0.006	19.0	-	0.036	0.025	0.049
residual error (additive)	0.016	0.007	-	-	0.015	0.000002	0.032

CL is clearance, V is volume of distribution, Q is inter-compartmental CL, IIV is inter-individual 555

variability, SE is standard error obtained with NONMEM 7.3 covariance step, CV is coefficient of 556

557 variation.

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Table 3: Summary of external evaluation with the evaluation dataset

Table 5. Sullil	nary or externa	ii evaii	iation	with the evalua	ation u	ataset			
dataset	Limit = 1 mg/L		Limit = 2 mg/L			PE (mg/L)	MPE	RMSE	
	n correct	OP	UP	n correct OP UP			(mg/L)	(mg/L)	
	(%)			(%)					
paired +	214/254	20	20	242/254	10	2	0.0004 (-1.07,	0.007	0.45
unpaired	(84.3)			(95.3)			0.84)		
paired:	53/57	3	1	56/57	1	0	-0.04 (-0.57,	-0.03	0.32
study≥1mg/L	(93.0)			(98.2)			0.70)		
paired:	31/33	2	0	33/33 (100)	0	0	-0.08 (-0.50,	-0.05	0.35
study≥2mg/L	(93.9)						0.74)		
paired:	19/20	0	1	20/20 (100)	0	0	-0.06 (-0.56,	-0.02	0.42
study≥3mg/L	(95.0)						0.82)		
unpaired	136/161	14	11	155/161	5	1	0.02 (-1.11,	-0.001	0.43
	(84.5)			(96.3)			0.70)		
XV: paired:	478/502	12	12	460/502	21	21	-0.04 (-1.77,	0.03	1.28
study≥3mg/L	(95.2)			(91.6)			3.03)		
XV: peaks ^a	-	-	-	-	-	-	0.16 (-4.76,	0.19	2.55
							5.01)		
AUC(0-t) ^a	-	-	-	-	-	-	10.8 (-24.9,	14.5 ^b	30.2 b
							62.2) ^b		

Correct indicates that the predicted trough concentration agrees with the measured concentration (is

562 above/below the limit); OP is overprediction, UP is underprediction; PE is prediction error (median

(95% confidence interval)), MPE is mean prediction error, RMSE is root mean square error, XV is 563

564 cross-validation. Except ^a all results refer to trough prediction evaluation. ^b in mg h/L.

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567	Figure legends
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569	Figure 1: Observed versus population predicted gentamicin serum concentrations (top left for the
570	model-building dataset and bottom left for the evaluation dataset) and conditional weighted residuals
571	versus time after dose (top right for the model-building dataset and bottom right for the evaluation
572	dataset).
573	
574	Figure 2: Visual predictive check of 1000 simulated concentration-time datasets from the final model,
575	using the model-building dataset (left) and the evaluation dataset (right). Points are the observations,
576	black lines are the 2.5^{th} , 50^{th} , and 97.5^{th} percentiles, and the shaded areas are the 95% confidence
577	intervals of the corresponding predicted gentamicin concentrations.
578	
579	Figure 3: Comparison of predictive performance of the developed model (shaded box plot) and
580	previously published neonatal gentamicin PK models.





