



## Comparison of associations between mortality and air pollution exposure estimated with a hybrid, a land-use regression and a dispersion model

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### ABSTRACT

**Introduction:** To characterize air pollution exposure at a fine spatial scale, different exposure assessment methods have been applied. Comparison of associations with health from different exposure methods are scarce. The aim of this study was to evaluate associations of air pollution based on hybrid, land-use regression (LUR) and dispersion models with natural cause and cause-specific mortality.

**Methods:** We followed a Dutch national cohort of approximately 10.5 million adults aged 29+ years from 2008 until 2012. We used Cox proportional hazard models with age as underlying time scale and adjusted for several potential individual and area-level socio-economic status confounders to evaluate associations of annual average residential NO<sub>2</sub>, PM<sub>2.5</sub> and BC exposure estimates based on two stochastic models (Dutch LUR, European-wide hybrid) and deterministic Dutch dispersion models.

**Results:** Spatial variability of PM<sub>2.5</sub> and BC exposure was smaller for LUR compared to hybrid and dispersion models. NO<sub>2</sub> exposure variability was similar for the three methods. Pearson correlations between hybrid, LUR and dispersion modeled NO<sub>2</sub> and BC ranged from 0.72 to 0.83; correlations for PM<sub>2.5</sub> were slightly lower (0.61–0.72). In general, all three models showed stronger associations of air pollutants with respiratory disease and lung cancer mortality than with natural cause and cardiovascular disease mortality. The strength of the associations differed between the three exposure models. Associations of air pollutants estimated by LUR were generally weaker compared to associations of air pollutants estimated by hybrid and dispersion models. For natural cause mortality, we found a hazard ratio (HR) of 1.030 (95% confidence interval (CI): 1.019, 1.041) per 10 µg/m<sup>3</sup> for hybrid modeled NO<sub>2</sub>, a HR of 1.003 (95% CI: 0.993, 1.013) per 10 µg/m<sup>3</sup> for LUR modeled NO<sub>2</sub> and a HR of 1.015 (95% CI: 1.005, 1.024) per 10 µg/m<sup>3</sup> for dispersion modeled NO<sub>2</sub>.

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**Conclusion:** Air pollution was positively associated with natural cause and cause-specific mortality, but the strength of the associations differed between the three exposure models. Our study documents that the selected exposure model may contribute to heterogeneity in effect estimates of associations between air pollution and health.

## 1. Introduction

A large number of epidemiological studies have shown associations of long-term exposure to ambient air pollution with mortality (Atkinson et al., 2018; Hoek et al., 2013). One of the main challenges of these studies is to assess residential long-term air pollution exposure at a fine spatial scale. Several exposure assessment methods have been developed and are now commonly applied (Jerrett et al., 2005; Hoek, 2017). Typically, studies have used land-use regression (LUR) or dispersion models to estimate long-term air pollution exposures (Jerrett et al., 2005; Hoek, 2017).

LUR and dispersion models are based on distinctly different methodological principles. LUR models are empirical models which use regression techniques to develop predictions based on air pollution measurements at a large number of sites and predictor data from geographic information systems (GIS). These empirical prediction models are then applied to non-measured locations. Dispersion models rely on deterministic equations and use data on emissions, source characteristics, chemical and physical properties of the pollutants, topography and meteorology modeling transport of pollutants through the atmosphere to estimate outdoor ground level air pollution concentrations (Jerrett et al., 2005; Hoek, 2017). Increased recognition of the limitations of both approaches has led to the development of hybrid models. These models combine data from dispersion models, land-use and surface monitoring data (Hoek, 2017) and are generally based on linear regression techniques, although Bayesian and machine learning methods (e.g. random forest) are increasingly used (Chen et al., 2019; Di et al., 2019; Cowie et al., 2019; Hanigan et al., 2017).

The use of different exposure assessment methods has been hypothesized to contribute to differences in effect estimates of associations between air pollution and health between epidemiological studies. Several studies showed moderate to good agreements between LUR and dispersion modeled air pollution concentrations (Dijkema et al., 2010; Marshall et al., 2008; Cyrus et al., 2005; Gulliver et al., 2011; Sellier et al., 2014; Hennig et al., 2016). De Hoogh et al. for example, reported a median correlation between LUR and dispersion model estimated nitrogen dioxide (NO<sub>2</sub>) concentrations of 0.75 and between LUR and dispersion Particulate Matter <2.5 µg (PM<sub>2.5</sub>) concentrations of 0.29 in 13 ESCAPE study areas (de Hoogh et al., 2014). However, at present, comparisons of effect estimates of different exposure assessment methods are scarce (Sellier et al., 2014; Wang et al., 2015; Jerrett et al., 2016).

This study is part of the ELAPSE project (*Effects of Low-Level Air Pollution: A Study in Europe*). Within the ELAPSE project, we evaluated associations of annual average air pollution concentrations based on a Europe-wide hybrid model including satellite observations, land-use predictors and dispersion model estimates in 11 cohorts with in-depth individual data and 7 large administrative/national cohorts (<http://www.elapseproject.eu/>). The aim of this study was to evaluate associations of long-term air pollution exposure based on hybrid, LUR and dispersion models with natural cause and cause-specific mortality in the Dutch national cohort (n ~ 10.5 million).

## 2. Methods

### 2.1. Study population and mortality outcomes

We created an administrative cohort that includes the full Dutch population aged 29+, on January 1, 2008, resulting in a study

population of approximately 10.5 million adults. The cohort was compiled based on data from several databases from Statistics Netherlands (CBS), including mortality and individual characteristics (such as sex, marital status, region of origin and standardized household income), as described elsewhere (Fischer et al., 2015; Fischer et al., 2020). We followed the cohort from 1 January 2008 until 31 December 2012.

We linked area-level socio-economic status (SES) indicators to the cohort to adjust for potential confounding possibly not accounted for by the available individual (SES) indicators. As SES has multiple dimensions (e.g. income, occupation) and correlations between air pollutants and each of these dimension may be different, we included several SES indicators. The following indicators were linked: mean income (mean income per income recipient), percentage of non-western immigrants and unemployment rate (number of people with income support per 1000 inhabitants aged 15–64 years) in 2006 at both regional (NUTS 3, n = 40) and neighborhood level (n ~ 2600, representing on average approximately 2900 addresses). NUTS (Nomenclature des Unités Territoriales Statistique) is a geocode standard for referencing the subdivisions of countries for statistical purposes and is developed and regulated by Eurostat, the statistical office of the European Union. Further, we used a composite SES score that represents the education, occupational and economic status at regional level and at four digit postal code level (PC4, n ~ 4000, representing on average approximately 1800 addresses). This composite score was only available at PC4 level; hence we do not have the composite score at neighborhood level.

As mortality outcomes, we selected natural cause (International Classification of Diseases, 10th Revision (ICD-10) codes: A00-R99), cardiovascular disease (I10-I70), respiratory disease (J00-J99) and lung cancer mortality (C34). Secondary analyses were conducted with more specific mortality outcomes: ischemic heart disease mortality (I20-I25; IHD), cerebrovascular mortality (I60-I69, CBV) and COPD mortality (J40-J44, J47).

### 2.2. Exposure assessment

We used annual average residential air pollution exposure estimates based on LUR models (Eeftens et al., 2012; Beelen et al., 2013), dispersion models (Keuken et al., 2013; Velders and Diederiksen, 2009) and hybrid models (De Hoogh et al., 2018), referred to as LUR, dispersion and hybrid models, respectively. Annual average residential NO<sub>2</sub> and PM<sub>2.5</sub> exposures were estimated by all three models. Black carbon (BC), measured as PM<sub>2.5</sub> absorbance based on reflectance measurement of the filters, was estimated by the LUR and hybrid model. Elemental carbon (EC) was estimated by the dispersion model. EC is often used as a proxy for BC and on average 1 unit PM<sub>2.5</sub> absorbance corresponds to 1.1 µg/m<sup>3</sup> elemental carbon (Janssen et al., 2011). A description of the different exposure assessment models is given below. Table 1 presents key features of the three methods, which are discussed in more detail below. The spatial variation of hybrid, LUR and dispersion modeled NO<sub>2</sub>, PM<sub>2.5</sub> and BC is shown in Supplemental Fig. S1.

#### 2.2.1. LUR model

LUR models were developed within the ESCAPE (European Study of Cohorts for Air Pollution Effects) project, using air pollution measurements collected during 2009 and 2010 (Eeftens et al., 2012; Beelen et al., 2013). Measurements were conducted at regional background, urban background and traffic sites throughout the Netherlands. Three 2-week measurements were conducted at 40 (PM<sub>2.5</sub> and BC) and 80 (NO<sub>2</sub>) sites

[of which 6 (PM<sub>2.5</sub> and BC) and 12 (NO<sub>2</sub>) were located in Belgium]. For each measurement site, results from the three 2-week measurements were averaged to estimate the annual average, adjusting for temporal variation (Eeftens et al., 2012; Beelen et al., 2013). Next, road length (Eurostreets digital road network, version 3.1 derived from the TeleAtlas MultiNet data set), traffic intensity (Nationaal WegenBestand) and European CORINE land cover predictor variables were calculated for each site, using the site coordinates and data within a GIS.

For NO<sub>2</sub>, PM<sub>2.5</sub> and BC, linear LUR models were developed using a supervised stepwise selection procedure (Eeftens et al., 2012; Beelen et al., 2013). For all available potential predictor variables, univariate linear regressions with annual average air pollution concentrations were evaluated first. The predictor variable with the highest adjusted explained variance (adjusted R<sup>2</sup>) was included in the LUR model if the direction of effect was as defined *a priori*. Next, the effect of the remaining predictor variables on the model adjusted R<sup>2</sup> was evaluated and the predictor variable giving the highest gain in adjusted R<sup>2</sup> and the right direction of effect was included in the model. This process continued until there were no more variables with the right direction of effect, which added at least 0.01 (1%) to the adjusted R<sup>2</sup> of the previous model. Details of the measurements and LUR modelling have been published elsewhere (Eeftens et al., 2012; Beelen, 2013).

The NO<sub>2</sub>, PM<sub>2.5</sub> and BC LUR models included several predictors, mainly traffic related indicators. Model performance was moderate to high; the explained variance was 86% for NO<sub>2</sub>, 67% for PM<sub>2.5</sub> and 92% for PM<sub>2.5</sub>abs (Eeftens et al., 2012; Beelen, 2013). LUR models are shown in Supplemental Table S1. The LUR models were used to estimate air pollution concentrations at each address in the Netherlands. LUR modeled NO<sub>2</sub> concentrations higher than 80 µg/m<sup>3</sup> (n ~ 300) were set to 80 µg/m<sup>3</sup> as these values are probably due to an unrealistic combination of explanatory variables (the maximum annual average NO<sub>2</sub> concentrations measured within the ESCAPE study was 61.5 µg/m<sup>3</sup>).

### 2.2.2. Dispersion model

The Dutch Operational Priority Substances (OPS) dispersion models are a combination of a Lagrangian trajectory model (for long-distance transport) and a Gaussian plume model (for the local scale road traffic contribution) (Van Jaarsveld, 2004). The Lagrangian trajectory model simulates atmospheric processes and estimates annual background concentration for NO<sub>2</sub> and PM<sub>2.5</sub> at a spatial resolution of 1 \* 1 km based on emission inventory data and meteorological parameters (Velders and Diederens, 2009). EC emissions are not included in the emission inventory for the Netherlands and European sources. Hence, fractions of EC in primary PM<sub>2.5</sub> emissions for all relevant sources, as developed in the EUCAARI (European Integrated project on Aerosol, Cloud, Climate, and Air Quality Interactions) European research project (<http://www.atm.helsinki.fi/eucaari/>), were used to estimate EC emission factors.

To estimate the contribution of road-traffic emissions to the 1\*1 km annual average background concentrations predicted by the Lagrangian trajectory model, two standard Dutch models were used (Keuken et al., 2013). For the contribution of road-traffic emissions inner urban roads, the street canyon model SRM1 (*Standaardrekenmethode 1*), is used. This model originates from the CAR model (Calculation of Air Pollution from

Road Traffic) and specifies a source–receptor as a function of the distance to the street axis for five different road types (Eerens et al., 1993). The contribution of traffic emissions to annual average concentrations depends on the emission rate, annual average wind speed, road type and distance to the road. For the contribution of road-traffic emissions from motorways, the line-source model SRM2 is used. This model is based on a Gaussian plume model which assumes that the contribution to ambient air pollution concentrations downwind of the motorway is dependent of the emission rate and the wind speed. Several factors, such as vehicle-induced turbulence, the upwind roughness of the terrain, the presence of noise screens near the motorway are taken into account in SRM2. Both models estimate the contribution of road-traffic air pollution on address level.

We used the average 2007–2009 OPS dispersion model background estimates. For the local scale road-traffic contribution, input data for the years 2007–2009 was not complete and reliable. Therefore, input data for the year 2014 was rescaled to the year 2008. More information about the dispersion models can be found elsewhere (Keuken et al., 2013; Velders and Diederens, 2009; Wesseling and Visser, 2003; Velders et al., 2020; Wesseling et al., 2011).

### 2.2.3. Hybrid model

Hybrid models were developed using measured NO<sub>2</sub> and PM<sub>2.5</sub> daily concentrations (aggregated to annual mean) for 2010, derived from the AirBase v8 dataset (EEA, 2015). BC concentrations were not available through AirBase, therefore, the ESCAPE annual mean BC concentrations based on measurements in Western-Europe, in the period 2009–2010, was used (De Hoogh et al., 2018). Potential predictor variables were prepared as a series of 100 \* 100 m rasters. Satellite derived NO<sub>2</sub> and PM<sub>2.5</sub> (~10 km resolution) were offered to the model (De Hoogh et al., 2018). In addition, concentrations estimates for 2010 from long-range chemical transport models (CTM) were offered as predictors (De Hoogh et al., 2018). Road length (Eurostreets digital road network, version 3.1 derived from the TeleAtlas MultiNet data set) and European CORINE land cover predictor variables were calculated for each raster. The integration of large scale satellite and CTM predictors distinguish the hybrid model from the classical LUR model.

Hybrid models were developed in a two-stage statistical procedure. First, LUR models were developed according to the ESCAPE protocol (used to develop the LUR models used in this paper), involving supervised stepwise linear regression (De Hoogh et al., 2018). Second, ordinary kriging was applied to the residuals of the model (De Hoogh et al., 2018). If kriging was not successful longitude and/or latitude were offered as additional predictors. The rationale for offering longitude and latitude was to explain large scale spatial trends in air pollution across Europe, that were not accounted for by the land use, traffic, CTM and satellite predictor variables. Annual average air pollution concentrations were estimated at a 100 \* 100 m spatial resolution.

The NO<sub>2</sub> hybrid model includes CTM NO<sub>2</sub>, roads, major roads, natural and residential predictor variables (De Hoogh et al., 2018). For NO<sub>2</sub>, no model was possible with both CTM and satellite estimates, as CTM and satellite estimates were moderately strongly correlated. We previously suggested that CTM's were better developed for NO<sub>2</sub> than for

**Table 1**  
Exposure assessment methods for long-term air pollution exposures.

Exposure assessment method	Monitoring approach	Modelling Approach	Area	Spatial scale	Year
LUR model	3 × 14-day average per year, 40–80 sites	Empirical (LUR)	Netherlands + Belgium	residential address	2009
Dispersion model	Dutch routine monitoring (45–60 sites)	Deterministic model	Netherlands	residential address	2008 <sup>a</sup>
Hybrid model	Annual average 436–2399 routine monitoring sites	Empirical (hybrid LUR with CTM, SAT <sup>b</sup> )	Western Europe	100 × 100 m grid	2010

<sup>a</sup> Based on average 2007–2009 Lagrangian trajectory model background estimates and local scale road-traffic contribution rescaled from 2014 to 2008.

<sup>b</sup> CTM = chemical transport model; SAT = satellite.

PM<sub>2.5</sub> when discussing the contribution of CTM and satellite estimates to PM<sub>2.5</sub> and NO<sub>2</sub> hybrid models (De Hoogh et al., 2018). The PM<sub>2.5</sub> hybrid model includes satellite and CTM PM<sub>2.5</sub> estimates, altitude, all roads, natural areas, ports and residential area. Kriging was applied to explain the left over residual variation (De Hoogh et al., 2018). The BC hybrid model includes CTM PM<sub>2.5</sub> estimates, roads, PM<sub>2.5</sub> satellite estimates, urban green, residential and natural land variables and latitude. The explained variance was 59% for NO<sub>2</sub>, 72% for PM<sub>2.5</sub> and 54% for BC (De Hoogh et al., 2018).

### 2.3. Statistical analyses

To evaluate associations of the air pollutants with natural cause and cause-specific mortality, we used Cox proportional hazard models. We specified *a priori* Cox models stratified by sex, with age as underlying time scale and increasing degrees of covariate adjustment. All models applied a correction of the standard error for clustering in neighborhoods. Model 1 included no covariates. Model 2 included individual-level covariate data on standardized household income, region of origin and marital status. Model 3 (main model) additionally included area level data on mean income per income recipient, unemployment rate, percentage non-western immigrants and the socio-economic composite score (the educational, occupational and economical status). In addition, the difference between neighborhood and the region of mean income, unemployment rate percentage, non-western immigrants and the socio-economic composite score were included.

We explored exposure-response curves for all exposure-mortality associations (main model) using natural splines with 3 degrees of freedom. Exposure-response curves were consistent with linearity along the most commonly observed range of the concentration or indicated supra-linear shapes with steeper slopes at the lower end (Fig. S2a–d). Hence, we decided to report results for linear exposure terms. Associations were expressed per 10 µg/m<sup>3</sup> for NO<sub>2</sub>, per 5 µg/m<sup>3</sup> for PM<sub>2.5</sub> and per 0.5 \* 10<sup>-5</sup>/m for BC (0.5 µg/m<sup>3</sup> for EC), based upon increments used within the ESCAPE project.

To evaluate potential mutual confounding of NO<sub>2</sub>, PM<sub>2.5</sub> and BC assessed by the same method, we specified two-pollutant models. Further, we evaluated a joint hazard ratio (JHR) of air pollutants to assess the joint risk of exposure to a mixture of these pollutants. The JHR can be assessed using the Cumulative Risk Index (CRI) method (Jerrett, 2013; Crouse et al., 2015). The JHR represents the hazard for an inter-quartile range (IQR) increase in both air pollutants relative to the odds for no increase in any of the pollutants.

We denote the JHR based on the combination of the P pollutants evaluated at x as the Cumulative Risk Index (CRI) which was defined as:

$$CRI = \exp \left\{ \sum_{p=1}^P \widehat{\beta}_p x_p \right\} = \exp(\widehat{\beta}'x) = \prod_{p=1}^P JHR_p$$

where  $\widehat{\beta}' = (\widehat{\beta}_1, \dots, \widehat{\beta}_p)$  are the estimates of the log hazard ratio for the P exposures estimated in a Cox proportional hazard model consisting of all P exposures together,  $x' = (x_1, \dots, x_p)$  are the levels at which each exposure-specific HR is evaluated and  $JHR_p = \exp(\widehat{\beta}_p x_p)$  denotes the JHR for the p<sup>th</sup> exposure in a multi-exposure model. JHRs were estimated assuming additive effect estimates (log hazard ratios) of the exposures. The 95% confidence interval of CRI is defined by:  $\exp\{\widehat{\beta}'x \pm 1.96 \times \sqrt{x' \times Cov(\widehat{\beta}) \times x}\}$ . This definition of the confidence interval is similar to that described elsewhere (Jerrett, 2013; Crouse et al., 2015).

All statistical analyses were conducted in R (<https://www.R-project.org/>), version 3.4.0, following centrally developed analysis scripts.

## 3. Results

### 3.1. Study population and mortality

Our cohort consisted of 10,532,360 subjects aged 29 year or older who contributed to 50,707,159 person-years follow-up. The mean age at baseline was 52.5 years, the majority of the subjects were married and from Dutch origin (Table 2). We observed 606,527 natural cause deaths, 165,601 cardiovascular disease deaths, 63,285 respiratory disease deaths and 49,488 lung cancer deaths. Of all cardiovascular deaths, ~30% died from IHD and ~25% died from CBV. Of all respiratory deaths, ~47% died because of COPD.

### 3.2. Exposure distribution and mutual correlations

Hybrid modeled NO<sub>2</sub> and BC concentrations were higher than LUR and dispersion modeled concentrations (Fig. 1). The spatial variation in NO<sub>2</sub> was quite similar across all three exposure models. The exposure range (5th – 95th percentile) was larger for dispersion modeled BC than for hybrid and LUR modeled BC. Mean PM<sub>2.5</sub> concentrations were quite similar between the different models, but the variation differed. The IQR of PM<sub>2.5</sub> estimated by the LUR model (0.87 µg/m<sup>3</sup>) was substantially lower than the IQR of PM<sub>2.5</sub> estimated by the hybrid model (1.90 µg/m<sup>3</sup>) and the dispersion model (2.15 µg/m<sup>3</sup>).

Pearson correlations between hybrid, LUR and dispersion modeled NO<sub>2</sub> at the residential addresses ranged from r = 0.78 to 0.83 (Table 3). Hybrid, LUR and dispersion modeled BC were slightly weaker correlated (Pearson r = 0.72–0.80). We found the lowest correlation between hybrid and LUR modeled PM<sub>2.5</sub> (Pearson r = 0.61). In general, NO<sub>2</sub> and BC from the same method were strongly correlated, while correlations of both exposures with PM<sub>2.5</sub> were lower (Table 3). Correlations between dispersion modeled NO<sub>2</sub>, PM<sub>2.5</sub> and BC were substantially stronger than correlations based on hybrid or LUR models.

### 3.3. Associations of air pollution with mortality

#### 3.3.1. Associations in single pollutants models

We found significant associations of air pollution with mortality outcomes in our main model (Table 4). For all three exposure models, the strongest associations were found with lung cancer mortality and the weakest associations with cardiovascular disease mortality, however the strength of the associations differed between the models. Across outcomes, effect estimates and statistical significance were more comparable between the hybrid and dispersion models. The LUR model generally showed weaker associations, except for lung cancer.

For natural cause mortality, we found associations for all three pollutants with the hybrid and dispersion model and no association with the LUR model. HRs were larger for the hybrid model for NO<sub>2</sub> and BC, but smaller for PM<sub>2.5</sub> compared to the dispersion model. For cardiovascular disease mortality, we found positive associations with the hybrid modeled pollutants and PM<sub>2.5</sub> from the dispersion model and no association with the LUR modeled pollutants. For respiratory disease mortality, we found the strongest associations with the hybrid and dispersion models for all three pollutants and no association with the LUR modeled pollutants. For lung cancer mortality, we found strong associations with all three models with similar HRs for all three pollutants.

In models without adjustments for individual and area-level SES indicators (Model 1), we found significant associations for all hybrid, LUR and dispersion modeled exposures and all outcomes, except for cardiovascular disease mortality (Fig. 2). The difference in strength of the associations in the minimally adjusted models was generally smaller between hybrid, LUR and dispersion modeled air pollution than in our main (fully adjusted) models, except for lung cancer mortality. For example, for natural cause mortality, the HR for LUR modeled PM<sub>2.5</sub> in the minimally adjusted model 1 was mildly larger compared to

**Table 2**  
Population characteristics (n = 10,532,360).

Covariate	Category	N (%) or mean (sd)
<i>Individual covariates</i>		
Age		52.5 (15.1)
Sex	male	5,129,824 (48.7)
	female	5,402,536 (51.3)
Marital status	married	6,633,882 (63.0)
	widowed	836,538 (7.9)
	divorced	1,058,624 (10.1)
	single	2,003,316 (19.0)
Region of origin	Morocco	138,078 (1.3)
	Turkey	171,837 (1.6)
	Suriname	183,153 (1.7)
	Antilles Netherlands non-western	58,971 (0.6)
		265,125 (2.5)
	western	1,009,761 (9.6)
	Dutch	8,705,435 (82.7)
Standardized household income	<1%	71,732 (0.7)
	1–5%	156,883 (1.5)
	5–10%	334,400 (3.2)
	10–25%	1,301,197 (12.4)
	25–50%	2,570,225 (24.5)
	50–75%	2,870,947 (27.3)
	75–90%	1,881,113 (17.9)
	90–95%	653,661 (6.2)
	95–99%	529,543 (5.0)
	>99%	133,457 (1.3)
<i>Area-level covariates</i>		
Composite SES 4 digit postal code	Based on education, income and paid occupation (year = 2007–2010)	0.02 (0.98)
Mean income neighborhood	Mean income per income recipient *€ 1000 (year = 2006)	18.23 (2.50)
Unemployment rate neighborhood	Number of people with income support per 1000 inhabitants of 15–64 years (year = 2006)	27.15 (8.78)
Percentage non-western immigrants neighborhood	Percentage non-western immigrants (year = 2006)	10.09 (11.87)
Composite SES region	Based on education, income and paid occupation (year = 2007–2010)	0.01 (0.27)
Mean income region	Mean income per income recipient *€ 1000 (year = 2006)	18.17 (1.16)
Unemployment rate region	Number of people with income support per 1000 inhabitants of 15–64 years (year = 2006)	27.35 (5.83)
Percentage non-western immigrants region	Percentage non-western immigrants (year = 2006)	10.47 (7.00)
<i>Mortality outcomes</i>		
natural cause mortality		606,527
cardiovascular mortality		165,601 49,248

**Table 2 (continued)**

Covariate	Category	N (%) or mean (sd)
ischemic heart disease mortality		
		40,597
cerebrovascular disease mortality		
		63,285
respiratory disease mortality		
		29,882
COPD mortality		49,488
Lung cancer mortality		

dispersion modeled PM<sub>2.5</sub>, whereas the HR in the fully adjusted model 3 was lower than for dispersion modeled PM<sub>2.5</sub>. For NO<sub>2</sub> and BC the same pattern was found for natural cause mortality. Further, we note that for all mortality outcomes, associations in the minimally adjusted model were strongest for hybrid modeled NO<sub>2</sub> and BC. Upon adjustment for potential confounders, hybrid modeled pollutants showed similar decreases in HRs compared to the LUR modeled pollutants, but the HRs in model 1 were generally larger for hybrid modeled pollutants than for LUR modeled pollutants. As a result, in the main model, associations of hybrid modeled pollutants with most outcomes remained significant.

We found weak associations of hybrid and dispersion modeled air pollution with CBV mortality and no associations with IHD mortality (Table S2). Hybrid and dispersion modeled air pollution were both positively associated with COPD mortality. LUR modeled pollutants were not significantly associated with IHD, CBV and COPD mortality.

### 3.3.2. Associations in multi pollutants models

In two pollutant models with combinations of hybrid modeled pollutants, associations of NO<sub>2</sub> and BC were barely affected by adjustment for PM<sub>2.5</sub>, except for lung cancer mortality (Table 5 for main mortality outcomes and Table S3 for secondary mortality outcomes). Associations of hybrid PM<sub>2.5</sub> with natural cause and respiratory disease mortality on the other hand attenuated and lost significance. In two pollutant models with combinations of dispersion modeled pollutants, associations of PM<sub>2.5</sub> with all outcomes remained (borderline) significant after adjustment for NO<sub>2</sub> or BC. Associations of dispersion modeled NO<sub>2</sub> or BC with all outcomes attenuated and lost significance after adjustment for PM<sub>2.5</sub>. For example, for respiratory disease mortality, the HR of dispersion modeled NO<sub>2</sub> changed from 1.036 (95% CI: 1.015, 1.058) to 0.928 (95% CI: 0.894, 0.963) after adjustment for PM<sub>2.5</sub> and the HR of dispersion modeled PM<sub>2.5</sub> changed from 1.126 (95% CI: 1.087, 1.167) to 1.250 (95% CI: 1.174, 1.331) after adjustment for NO<sub>2</sub>.

For all three models, JHR of combinations of pollutants, expressed per IQR increase (not per fixed increment), were similar or slightly larger compared to HRs of single pollutant models expressed per IQR (Fig. S3). For example, for lung cancer mortality, the HR of hybrid modeled NO<sub>2</sub> was 1.084 (95% CI: 1.063, 1.105) per IQR increase and the JHR of hybrid modeled NO<sub>2</sub> and PM<sub>2.5</sub> was 1.090 (95% CI: 1.069, 1.112) per IQR increase and the HR of dispersion modeled PM<sub>2.5</sub> was 1.066 (95% CI: 1.047, 1.085) per IQR increase and the JHR of dispersion modeled PM<sub>2.5</sub> and NO<sub>2</sub> was 1.068 (95% CI: 1.049, 1.088). JHRs based on hybrid modeled pollutants were fairly similar in strength. JHR based on dispersion modeled pollutants were generally largest for a combination of PM<sub>2.5</sub> and BC.

## 4. Discussion

Exposure estimates from the hybrid, LUR and dispersion models were moderately to strongly correlated. We found generally positive associations of air pollution with natural cause, cardiovascular disease, respiratory disease and lung cancer mortality, but the strength of the associations differed between the three exposure models. For all three models, the strongest associations were found with lung cancer mortality and the weakest associations with cardiovascular disease

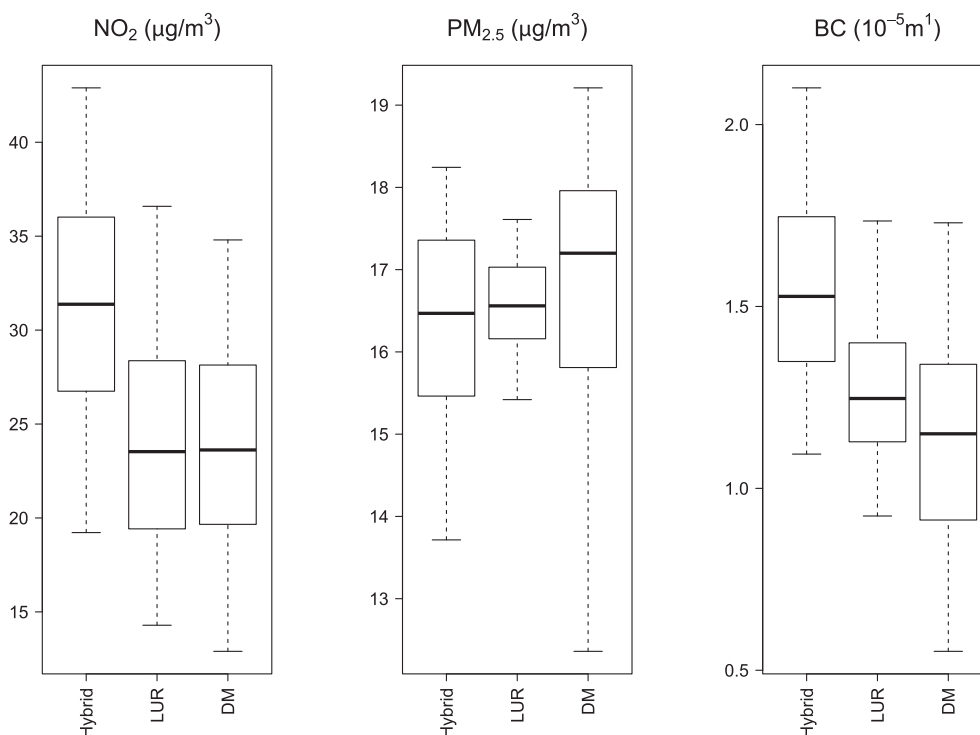


Fig. 1. Boxplots of NO<sub>2</sub>, PM<sub>2.5</sub> and BC concentrations based on a hybrid, LUR and dispersion (DM) model.<sup>a,b</sup> (<sup>a</sup>The middle box represents the middle 50% of the concentration, the line that divides the box in two is the median. The upper and lower end of the whiskers represent the 5th and 95th percentile. <sup>b</sup> DM BC is modeled per µg/m<sup>3</sup>).

Table 3  
Pearson correlations between NO<sub>2</sub>, PM<sub>2.5</sub> and BC based on a hybrid, LUR and dispersion (DM) model and mean concentrations (standard deviation = sd) of each pollutant.<sup>a</sup>

Pearson correlation		Hybrid			LUR			DM		
		NO <sub>2</sub>	PM <sub>2.5</sub>	BC	NO <sub>2</sub>	PM <sub>2.5</sub>	BC	NO <sub>2</sub>	PM <sub>2.5</sub>	BC
Hybrid	NO <sub>2</sub>	1.00								
	PM <sub>2.5</sub>	0.58	1.00							
	BC	0.88	0.52	1.00						
LUR	NO <sub>2</sub>	0.83	0.43	0.77	1.00					
	PM <sub>2.5</sub>	0.52	0.61	0.56	0.45	1.00				
	BC	0.78	0.44	0.80	0.86	0.74	1.00			
DM	NO <sub>2</sub>	0.79	0.56	0.74	0.78	0.63	0.79	1.00		
	PM <sub>2.5</sub>	0.71	0.69	0.66	0.66	0.72	0.70	0.86	1.00	
	BC	0.75	0.51	0.72	0.72	0.60	0.75	0.94	0.85	1.00
Mean (sd)		31.3 µg/m <sup>3</sup> (7.1)	16.3 µg/m <sup>3</sup> (1.4)	1.6 * 10 <sup>-5</sup> /m (0.3)	24.3 µg/m <sup>3</sup> (6.9)	16.6 µg/m <sup>3</sup> (0.7)	1.3 * 10 <sup>-5</sup> /m (0.2)	23.9 µg/m <sup>3</sup> (6.8)	16.7 µg/m <sup>3</sup> (2.0)	1.1 µg/m <sup>3</sup> (0.4)

<sup>a</sup> Mutual correlations are given in gray cells.

mortality. Air pollution modeled with the hybrid and dispersion models were generally significantly associated with all mortality outcomes, whereas air pollutants modeled with LUR were only significantly associated with lung cancer mortality. Differences between the three models

were smaller for the minimally adjusted confounder model than the main model, suggesting that sensitivity to confounding differed between the three exposure models. Two pollutant models suggested more robust associations with NO<sub>2</sub> for the hybrid model and with PM<sub>2.5</sub> for the

**Table 4**

Associations of air pollution based on a hybrid, LUR and dispersion (DM) model on natural cause, cardiovascular, respiratory and lung cancer mortality in single-pollutant models.<sup>a,b</sup>

Outcome	Pollutant	Hybrid	LUR	DM
		HR (95% CI)	HR (95% CI)	HR (95% CI)
Natural cause mortality	NO <sub>2</sub>	1.030	1.003	1.015
		(1.019, 1.041)	(0.993, 1.013)	(1.005, 1.024)
		1.021	1.006	1.035
	PM <sub>2.5</sub>	(0.999, 1.044)	(0.973, 1.040)	(1.018, 1.052)
		1.030	1.005	1.018
		(1.019, 1.041)	(0.993, 1.017)	(1.009, 1.027)
	BC	1.017	0.984	1.000
		(1.003, 1.031)	(0.970, 0.999)	(0.988, 1.013)
		1.015	0.994	1.021
Cardiovascular disease mortality	PM <sub>2.5</sub>	(0.988, 1.042)	(0.952, 1.039)	(0.998, 1.044)
		1.018	0.991	1.006
		(1.003, 1.033)	(0.975, 1.007)	(0.994, 1.018)
BC	1.038	0.996	1.036	
	(1.015, 1.062)	(0.972, 1.019)	(1.015, 1.058)	
	1.058	1.059	1.126	
Respiratory disease mortality	PM <sub>2.5</sub>	(1.009, 1.111)	(0.989, 1.133)	(1.087, 1.167)
		1.052	1.005	1.043
		(1.028, 1.076)	(0.981, 1.031)	(1.024, 1.062)
BC	1.091	1.089	1.079	
	(1.069, 1.114)	(1.065, 1.113)	(1.056, 1.102)	
	1.169	1.243	1.155	
Lung cancer mortality	PM <sub>2.5</sub>	(1.121, 1.218)	(1.155, 1.337)	(1.112, 1.198)
		1.085	1.087	1.071
		(1.062, 1.108)	(1.061, 1.114)	(1.051, 1.091)

<sup>a</sup> Associations are expressed per 10 µg/m<sup>3</sup> for NO<sub>2</sub>, per 5 µg/m<sup>3</sup> for PM<sub>2.5</sub> and per 0.5 \* 10<sup>-5</sup>/m for BC (0.5 µg/m<sup>3</sup> for EC).

<sup>b</sup> Associations of main model are adjusted for age, strata(sex), random (neighborhood), standardized household income, region of origin, marital status, socio-economic composite score region, mean income per income recipient region, unemployment rate region, percentage non-western immigrants region, and the difference between neighborhood and region of mean income, unemployment rate, non-western immigrants and the composite SES score (4 digit postal code).

dispersion model.

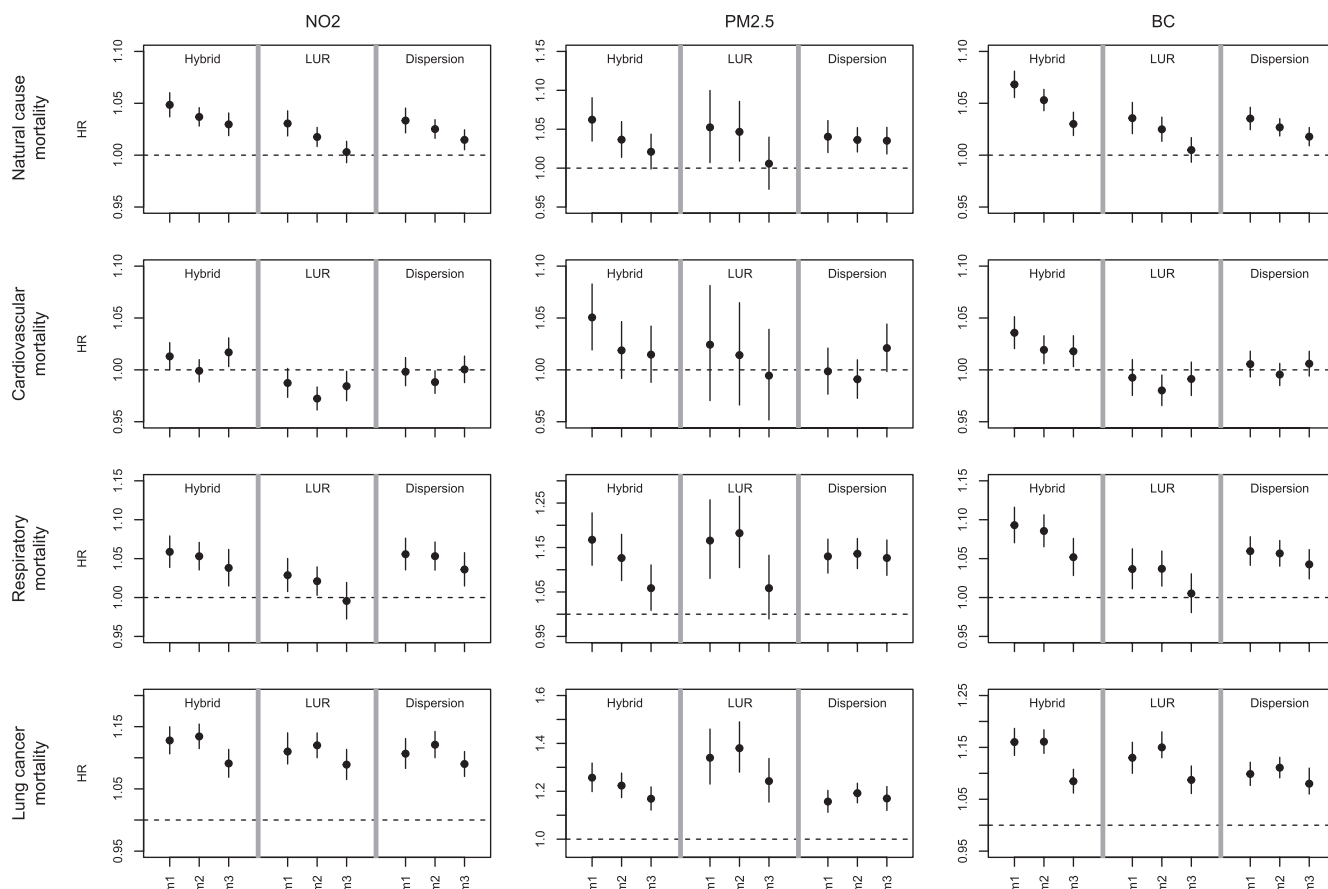
#### 4.1. Hybrid, LUR and dispersion modeled air pollution exposure patterns

Correlations between hybrid, LUR and dispersion modeled air pollution were strongest for NO<sub>2</sub>. This is likely due to the strong contribution of traffic to the small-scale variation of NO<sub>2</sub> within the Netherlands. The LUR and hybrid model differ from the dispersion model in that they do not rely on dispersion processes and assumptions about (traffic) emissions, but only use land use data (LUR model), such as traffic intensity, or a combination of land use data, satellite observations and dispersion model estimates (hybrid model). However, land use data, such as traffic intensity and population density, is used in the calculations of emissions and dispersions of air pollution in the dispersion models. Correlations between hybrid, LUR and dispersion modeled BC were slightly lower than for NO<sub>2</sub>, despite the fact that BC is also strongly determined by traffic. This might be due to the smaller number of BC measurements and a lack of emissions data. BC measurements are

not available through AirBase and BC emissions are not included in the emission inventory for the Netherlands. Instead, PM<sub>2.5</sub> satellite and CTM are included in the hybrid BC model, as indicators of BC emissions, which may have resulted in lower model performance. Because of the lack of BC emission data, assumed fractions of EC in primary PM<sub>2.5</sub> emissions for all relevant sources are used to estimate EC by the dispersion model. Correlations between hybrid, LUR and dispersion modeled PM<sub>2.5</sub> were lowest. This could be due to the low spatial variation of PM<sub>2.5</sub> within the Netherlands. The influence of traffic sources on PM<sub>2.5</sub> concentrations is lower and the influence of other sources such as industry is higher compared to NO<sub>2</sub> and BC concentrations (Eeftens et al., 2012).

De Hoogh et al. previously reported a strong correlation for LUR and dispersion modeled NO<sub>2</sub> (median Pearson correlation of 0.75) and a weak correlation for LUR and dispersion modeled PM<sub>2.5</sub> (median Pearson correlation of 0.29) within 13 ESCAPE study areas (de Hoogh et al., 2014). Correlations between dispersion modeled air pollution and measured air pollution at the Dutch ESCAPE measurement sites was 0.85 for NO<sub>2</sub> and 0.54 for PM<sub>2.5</sub> (de Hoogh et al., 2014). However, the correlations between the dispersion and the LUR model of approximately 1000 Dutch children of the Prevention and Incidence of Asthma and Mite Allergy birth cohort were strong for NO<sub>2</sub> and PM<sub>2.5</sub> (Pearson R = 0.89 for NO<sub>2</sub> and 0.81 for PM<sub>2.5</sub>) (Wang et al., 2015). Sellier et al. and Wang et al. also reported strong correlations between LUR and dispersion modeled NO<sub>2</sub> (Pearson R > 0.85) (Sellier et al., 2014; Wang et al., 2015). Wang et al. also reported a strong correlation between LUR and dispersion modeled PM<sub>2.5</sub> (Pearson R = 0.86) (Wang et al., 2015).

Hybrid modeled NO<sub>2</sub> concentrations were higher compared to LUR and dispersion modeled NO<sub>2</sub> concentrations. De Hoogh et al. (De Hoogh et al., 2018) reported an overestimation (fractional bias) of 13% for hybrid modeled NO<sub>2</sub> compared to ESCAPE NO<sub>2</sub> measurements in the overall ELAPSE area and a 26% overestimation for the Dutch ESCAPE NO<sub>2</sub> measurements. This overestimation is in line with the differences between hybrid and LUR modeled concentrations reported in this study. In the ESCAPE project, NO<sub>2</sub> was measured with Ogawa badges, which resulted in lower concentrations compared to chemiluminescence measurements on which the hybrid model was based (Cyrys et al., 2012). In the Netherlands, the Ogawa measurements were about 20% lower than the concurrent chemiluminescence measurements (Cyrys et al., 2012). We do not have an explanation why the hybrid model predicts higher concentrations than the dispersion model, which fits the Dutch monitoring data well. This may be a limitation of applying a European model in a single country. The models also differed in locations of monitoring sites, which may have influenced the concentrations levels. Monitoring sites used for LUR models were selected to represent residential exposure, i.e. sites were located near building facades representative for residential addresses. Monitoring stations used for dispersion and hybrid modelling on the other hand, were used for regulatory purposes and not all located near residential addresses. The higher hybrid modeled NO<sub>2</sub> concentrations (year = 2010) compared to LUR (year = 2009) and dispersion modeled NO<sub>2</sub> (year = 2008) are likely not due to the different year. NO<sub>2</sub> concentrations measured in the Dutch monitoring network showed that the average NO<sub>2</sub> concentrations in the Netherlands were somewhat lower in 2010 compared to 2009 and 2008 (RIVM, DCMR, and GGD Amsterdam, 2018). However, we note that LUR model estimates are based on 3 two-week measurements and recalibrated to annual averages, while hybrid (NO<sub>2</sub> and PM<sub>2.5</sub>) and dispersion model estimates are based on/validated with continuous measurements from monitoring stations. PM<sub>2.5</sub> concentrations were fairly similar between the three methods. However, LUR modeled PM<sub>2.5</sub> has a substantially lower spatial variation than hybrid and dispersion modeled PM<sub>2.5</sub>. This is likely due to the limited contrast in PM<sub>2.5</sub> concentrations measured at urban background and regional background in the Netherlands within the ESCAPE project (Eeftens et al., 2012). Further, large scale satellite, CTM predictors (hybrid) and the Lagrangian trajectory model (dispersion) may better capture long-range transport than



**Fig. 2.** Associations of air pollution based on a hybrid, LUR and dispersion model with natural cause, cardiovascular, respiratory and lung cancer mortality in models with increasing degree of adjustment for potential confounders.<sup>a,b</sup>. (<sup>a</sup>Associations are expressed per 10  $\mu\text{g}/\text{m}^3$  for  $\text{NO}_2$ , per 5  $\mu\text{g}/\text{m}^3$  for  $\text{PM}_{2.5}$  and per  $0.5 \cdot 10^{-5}/\text{m}$  for BC ( $0.5 \mu\text{g}/\text{m}^3$  for EC). Model 1 (m1) included the baseline hazard, a strata term for sex and a cluster for neighborhood. Model 2 (m2) additionally included standardized household income, region of origin and marital status. Model 3 (m3, main model) additionally included mean income per income recipient of the region, unemployment rate of the region, percentage non-western immigrants of the region and the socio-economic composite score (the educational, occupational and economical status) of the region, the difference between mean income per income recipient of the neighborhood and that of the region, the difference between unemployment rate of the neighborhood and that of the region, the difference between percentage non-western immigrants of the neighborhood and that of the region, the difference between the socio-economic composite score (based on the educational, occupational and economical status) at a four digit postal code level and that of the region).

by monitoring alone.

#### 4.2. Differences in effect estimates between hybrid, LUR and dispersion modeled air pollution

Based on the moderate to strong correlations between exposure estimates of the three models, we had expected more consistency in effect estimates of an association between air pollution and mortality. We do not know which of the three exposure methods performs best. Differences in effect estimates in the fully adjusted models could be explained by differences in exposure measurement error, in predicted exposure contrasts and in sensitivity to adjustment for confounders.

**Exposure measurement error** effects can lead to an attenuation of effect estimates and an increase of the confidence intervals (Basagaña et al., 2012). The impact of exposure measurement error on the estimated health effects is complex, as it depends on the combination of errors (Berkson, classical) (Samoli et al., 2020; Butland et al., 2020). The combination of errors may differ between pollutants and between methods. Performance of exposure assessment models are generally evaluated by  $R^2$  measures. However, published validation statistics of the hybrid and LUR models cannot be directly compared. Performance of hybrid models was based on (hold out validation of) measurement sites across Western-Europe, while validation for LUR models was based on leave-one-out cross validation of Dutch ESCAPE measurements sites.

Performance of the hybrid models for the Netherlands only, showed that the explained variation of the ESCAPE measurements of the hybrid  $\text{NO}_2$  model was high in the Netherlands ( $R^2 = 76\%$ , 80 sites), in contrast to the explained variation of the hybrid  $\text{PM}_{2.5}$  model ( $R^2 = 13\%$ , 40 sites) (De Hoogh et al., 2018). Further, we note that 40 measurement sites were used to develop the Dutch  $\text{PM}_{2.5}$  and BC LUR model (Beelen et al., 2013), while 543 measurement sites were used to develop the hybrid  $\text{PM}_{2.5}$  model and 436 sites were used to develop the hybrid BC model (De Hoogh et al., 2018). Models based on a small number of measurement sites tend to give higher  $R^2$  and leave-one out cross-validated  $R^2$  than those based on more sites (Basagaña et al., 2012). A larger number of sites will produce more stable models (Basagaña et al., 2012; Wang et al., 2013), which was the rationale to develop the Europe-wide hybrid models in the ELAPSE project.

National exposure assessment models, such as the LUR and dispersion models, may better capture national-specific small-scale variation patterns than the Western-Europe wide hybrid model, as relations between air pollution and predictor variables may differ between countries (De Hoogh et al., 2018). The hybrid model tends to average intra-study area differences in air pollution - predictor variables relations over entire study area (De Hoogh et al., 2018). However, as the LUR models used in this study are based on a relatively small number of measurement sites, they may not capture small-scale air pollution variations better than dispersion or hybrid models. Further, differences in air



**Table 5**

Associations of air pollution based on a hybrid, LUR and dispersion model with natural cause, cardiovascular, respiratory and lung cancer mortality in multi-pollutant models.<sup>a,b</sup>

Outcome	Pollutant	Adj. for PM <sub>2.5</sub>			Adj. for BC			Adj. for NO <sub>2</sub>		
		HR (95% CI)			HR (95% CI)			HR (95% CI)		
		Hybrid	LUR	DM	Hybrid	LUR	DM	Hybrid	LUR	DM
Natural cause mortality	NO <sub>2</sub>	1.039 (1.025, 1.054)	1.003 (0.991, 1.015)	0.992 (0.977, 1.008)	1.017 (0.997, 1.038)	0.999 (0.982, 1.016)	0.975 (0.950, 1.001)	.	.	.
	PM <sub>2.5</sub>	.	.	.	0.988 (0.963, 1.015)	0.982 (0.927, 1.042)	1.023 (0.996, 1.05)	0.972 (0.944, 1.000)	1.001 (0.963, 1.041)	1.046 (1.017, 1.076)
	BC	1.033 (1.020, 1.047)	1.010 (0.989, 1.031)	1.008 (0.994, 1.022)	.	.	.	1.015 (0.994, 1.036)	1.006 (0.986, 1.026)	1.039 (1.014, 1.065)
Cardio-vascular disease mortality	NO <sub>2</sub>	1.021 (1.002, 1.040)	0.980 (0.964, 0.997)	0.969 (0.948, 0.990)	1.007 (0.984, 1.032)	0.976 (0.954, 0.998)	0.960 (0.926, 0.995)	.	.	.
	PM <sub>2.5</sub>	.	.	.	0.996 (0.965, 1.028)	1.047 (0.967, 1.133)	1.036 (1.000, 1.073)	0.988 (0.953, 1.025)	1.027 (0.977, 1.08)	1.067 (1.028, 1.108)
	BC	1.019 (1.001, 1.037)	0.978 (0.950, 1.007)	0.991 (0.972, 1.009)	.	.	.	1.011 (0.985, 1.038)	1.013 (0.988, 1.038)	1.041 (1.006, 1.076)
Respiratory disease mortality	NO <sub>2</sub>	1.033 (1.001, 1.065)	0.982 (0.956, 1.010)	0.928 (0.894, 0.963)	0.983 (0.941, 1.027)	0.980 (0.943, 1.018)	0.946 (0.891, 1.004)	.	.	.
	PM <sub>2.5</sub>	.	.	.	1.010 (0.953, 1.070)	1.145 (1.013, 1.294)	1.176 (1.108, 1.249)	1.016 (0.952, 1.085)	1.090 (1.007, 1.178)	1.250 (1.174, 1.332)
	BC	1.049 (1.021, 1.078)	0.967 (0.925, 1.011)	0.972 (0.943, 1.003)	.	.	.	1.067 (1.021, 1.116)	1.023 (0.983, 1.065)	1.092 (1.035, 1.152)
Lung cancer mortality	NO <sub>2</sub>	1.063 (1.035, 1.092)	1.072 (1.046, 1.099)	1.018 (0.985, 1.052)	1.075 (1.035, 1.116)	1.074 (1.039, 1.110)	1.040 (0.980, 1.100)	.	.	.
	PM <sub>2.5</sub>	.	.	.	1.110 (1.058, 1.163)	1.061 (0.942, 1.195)	1.120 (1.060, 1.190)	1.082 (1.026, 1.141)	1.106 (1.021, 1.200)	1.126 (1.062, 1.194)
	BC	1.054 (1.028, 1.081)	1.069 (1.027, 1.112)	1.020 (0.990, 1.050)	.	.	.	1.018 (0.979, 1.059)	1.020 (0.984, 1.059)	1.040 (0.990, 1.090)

<sup>a</sup> Associations are expressed per 10 µg/m<sup>3</sup> for NO<sub>2</sub>, per 5 µg/m<sup>3</sup> for PM<sub>2.5</sub> and per 0.5 \* 10<sup>-5</sup>/m for BC (0.5 µg/m<sup>3</sup> for EC).

<sup>b</sup> Associations of main model are adjusted for age, strata(sex), random(neighborhood), standardized household income, region of origin, marital status, socio-economic composite score region, mean income per income recipient region, unemployment rate region, percentage non-western immigrants region, and the difference between neighborhood and region of mean income, unemployment rate, non-western immigrants and the composite SES score (4 digit postal code).

pollution - predictor variables relations in the Netherlands and in Western-Europe might be limited. Another notable difference is that the hybrid models estimated air pollution concentrations on a 100 \* 100 m grid while the LUR and dispersion modeled pollutants are estimated on address level. Address level estimates might be more accurate, but they may also be more sensitive to geocoding errors. We accounted for major geocoding errors in the LUR model (e.g. unreasonably small distance between address geocode and nearest road) by truncating the predictor variables to the minimum or maximum observed at the monitoring sites. Because the LUR, dispersion and hybrid models were independently developed over a different area and using different input data, our ability to attribute the differences in effect estimates to specific components of the exposure modelling method was somewhat hindered.

**Predicted concentration ranges** differed between hybrid, LUR and dispersion modeled pollutants. The IQR for LUR modeled BC and especially PM<sub>2.5</sub> were lower than for hybrid and dispersion modeled BC and PM<sub>2.5</sub>. HRs are expressed per fixed increment to be able to compare effects of hybrid, LUR and dispersion modeled pollutants. However, HRs of PM<sub>2.5</sub> are expressed per 5 µg/m<sup>3</sup> which is more than five times as large as the IQR of LUR modeled PM<sub>2.5</sub> and more than twice as large as the difference between the 95th – 5th percentile. A lower spatial variation limits the ability to capture differences in event rates across exposure

ranges. However, if differences in event rates are captured by LUR modeled PM<sub>2.5</sub>, HR expressed per 5 µg/m<sup>3</sup> can be very large, such as the association of LUR modeled PM<sub>2.5</sub> with lung cancer mortality. Effect estimates expressed per IQR were weaker for LUR modeled PM<sub>2.5</sub> than for hybrid and dispersion modeled PM<sub>2.5</sub> (Fig. S3).

We observed fairly similar effect estimates in the minimally adjusted models. In the fully adjusted models, where we adjusted for individual, neighborhood- and regional-level SES variables, the differences in effect estimates were larger. By adjusting for area-level SES variables, some of the neighborhood and regional scale variation in mortality was removed from the air pollution estimates. As the exposure assessment models differ in structure (i.e. importance of large- and small-scale predictors), they may be differentially sensitive to relations with area-level variables. The higher sensitivity of the hybrid and LUR models to **adjustment for confounders** may be due to the use of more generic predictor variables such as population density compared to specific air pollution emissions in the dispersion model. Further, we note that we did not adjust for traffic noise and personal lifestyle factors, such as smoking status and BMI, that may have impacted effect estimates of the models differently.

Correlations between air pollutants were much stronger in the dispersion model compared to the hybrid and LUR model. This may be

due to the inclusion of different predictors and zones of influence (e.g. size of circular buffers) in the hybrid and LUR models versus the use of the same fundamental dispersion processes and assumptions about emissions for all pollutants in the dispersion models (Fecht et al., 2016). The much stronger correlation between PM<sub>2.5</sub> and NO<sub>2</sub>/BC in the dispersion model could be due to an overestimation of the contribution of traffic to the total PM<sub>2.5</sub> concentration in the dispersion model or an underestimation of the contribution of traffic to the total PM<sub>2.5</sub> concentration in the hybrid and LUR model. We found that dispersion modeled PM<sub>2.5</sub> was slightly stronger correlated with hybrid modeled NO<sub>2</sub> (Pearson  $r = 0.71$ ) than with hybrid modeled PM<sub>2.5</sub> (Pearson  $r = 0.69$ ). Compared to hybrid and LUR modeled pollutants, the high correlation between dispersion modeled pollutants resulted in unstable effect estimates. In two-pollutant models with combinations of dispersion modeled pollutants, PM<sub>2.5</sub> remained associated with natural cause, cardiovascular and respiratory disease mortality outcomes after adjustment for NO<sub>2</sub> or BC. The opposite pattern was seen for hybrid modeled pollutants. Despite the weaker correlations of hybrid and LUR pollutants, JHRs showed a similar pattern for all three methods. The JHRs of the air pollutant mixture were similar or only slightly higher compared to HRs from single pollutant models for each exposure assessment methods. This suggests that a single pollutant could be sufficient to characterize the toxicity of the air pollution mixture (of NO<sub>2</sub>, PM<sub>2.5</sub> and BC).

#### 4.3. Comparison with previous studies

There are only a few studies that compared effect estimates of LUR and dispersion models. Sellier et al. (2014) previously compared associations of LUR and dispersion modeled NO<sub>2</sub> with birth weight. Most associations were weak and not significant and differed only slightly between the exposure models (Sellier et al., 2014). Wang et al. (2015) reported similar significant associations of LUR and dispersion modeled NO<sub>2</sub> and BC with lung function (FEV<sub>1</sub>, FVC) in Dutch children, while effect estimates (expressed per 5 µg/m<sup>3</sup>) with LUR modeled PM<sub>2.5</sub> were two times higher than for dispersion modeled PM<sub>2.5</sub>. This could be due to the difference in exposure contrast; the IQR for dispersion modeled PM<sub>2.5</sub> (3.7 µg/m<sup>3</sup>) was three times larger than the IQR for LUR modeled PM<sub>2.5</sub> (1.1 µg/m<sup>3</sup>) (Wang et al., 2015).

Associations of hybrid and dispersion NO<sub>2</sub> with natural cause, cardiovascular and respiratory disease mortality in our study were in line with recent HRs from meta-analysis (Atkinson et al., 2018). Associations of hybrid and dispersion PM<sub>2.5</sub> were slightly weaker than HRs from meta-analysis for cardiovascular disease mortality and stronger than HRs from meta-analysis for respiratory disease mortality (Hoek et al., 2013). We found the strongest associations with lung cancer mortality, but note that we lack information about smoking. However, a study by Fischer et al. (Fischer et al., 2015) suggested that it is unlikely that uncontrolled confounding from smoking or BMI substantially biased associations of air pollution with mortality in the Dutch national cohort. Klompaker et al. (2020) found no associations of air pollutants, modeled by the same LUR model as used in this article, with natural cause and cause-specific mortality in a 5 year follow-up of a large national health survey in models with adjustment for individual SES indicators and lifestyle factors and in models with adjustment for individual, neighborhood and regional SES. The authors speculated that the null findings might be due to the short follow-up period (Klompaker et al., 2020).

Two previously published studies reported positive associations of exposure to air pollution with mortality in a Dutch administrative cohort (including > 7 million individuals aged 30 years or older) (Fischer et al., 2015, 2020). Fischer et al. (2015) used different LUR models than we used. Associations of their LUR modeled NO<sub>2</sub> were different from associations of our LUR modeled NO<sub>2</sub>, but similar to associations of hybrid and dispersion modeled NO<sub>2</sub>. Fischer et al. (2020) used the same dispersion model as we used and reported slightly stronger associations

for PM<sub>2.5</sub> and BC with natural cause mortality. We note some important differences, as both studies performed by Fischer and colleagues excluded all individuals who moved 5 years before the start of the follow-up period (Fischer et al., 2015; Fischer et al., 2020). This may have contributed to slightly stronger associations compared to the associations in this study. In addition, the follow-up period of Fischer et al. (2015) and of Fischer et al. (2020) was from 2004 to 2011 and from 2008 to 2015, respectively. Furthermore, both studies only adjusted for neighborhood SES composite score as area-level SES indicator, while we included several neighborhood and regional level SES indicators. Associations of NO<sub>2</sub> with natural cause mortality with limited adjustments for neighborhood and regional SES were similar to associations of our main model. For example, for natural cause mortality, for models with adjustments for only regional and neighborhood composite SES, we found a HR of 1.022 (95% CI: 1.014, 1.030) for hybrid modeled NO<sub>2</sub>, a HR of 1.005 (95% CI: 0.994, 1.016) for LUR modeled NO<sub>2</sub> and a HR of 1.014 (95% CI: 1.006, 1.022) for dispersion modeled NO<sub>2</sub>. For models with adjustment for neighborhood and regional mean income, percentage of non-western immigrants and unemployment rate, but not for composite SES, we found a HR of 1.032 (95% CI: 1.021, 1.043) for hybrid modeled NO<sub>2</sub>, a HR of 1.003 (95% CI: 0.999, 1.007) for LUR modeled NO<sub>2</sub> and a HR of 1.016 (95% CI: 1.007, 1.026) for dispersion modeled NO<sub>2</sub>. We acknowledge that some over-adjustment is possible, but preferred this to inadequate adjustment for SES.

#### 5. Conclusion

Air pollution exposure estimates from a hybrid model, LUR and dispersion model were moderately to strongly correlated. We found generally positive associations of air pollution with natural cause, cardiovascular disease, cerebrovascular disease, respiratory disease, COPD and lung cancer mortality, but not with ischemic heart disease mortality. Despite the strong mutual correlations, the strength of the associations differed between the three exposure models. For all three models, the strongest associations were found with lung cancer mortality and the weakest associations with cardiovascular disease mortality. Air pollution modeled by the hybrid and dispersion models were generally more strongly associated with mortality than air pollution modeled by the LUR. Two pollutant models suggested more robust associations with NO<sub>2</sub> for the hybrid model and with PM<sub>2.5</sub> for the dispersion model. The difference between effect estimates depended on the mortality outcome. Differences in effect estimates between models are likely due to different measurement error, different sensitivity to confounding and different predicted exposure contrasts. Overall, our study documents that the selected exposure model may contribute to heterogeneity in effect estimates from cohort studies of long-term exposure to outdoor air pollution and mortality.

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#### CRediT authorship contribution statement

**Jochem O. Klompaker:** Formal analysis, Methodology, Software, Visualization, Writing - original draft. **Nicole Janssen:** Data curation, Software, Supervision, Writing - review & editing. **Zorana J. Andersen:** Writing - review & editing. **Richard Atkinson:** Writing - review & editing. **Mariska Bauwelinck:** Writing - review & editing. **Jie Chen:** Project administration, Writing - review & editing. **Kees de Hoogh:** Methodology, Software, Writing - review & editing. **Danny Houthuijs:** Writing - review & editing. **Klea Katsouyanni:** Methodology, Writing -

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## Declaration of Competing Interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

## Appendix A. Supplementary material

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.envint.2020.106306>.

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